

**Peyresq 2019 - Riemannian optimization - PA Absil**

30 June – 6 July 2019  
Project

## 1 Context

The cocktail party problem consists in recovering statistically independent source signals from linear combinations of these sources (see figure 1).

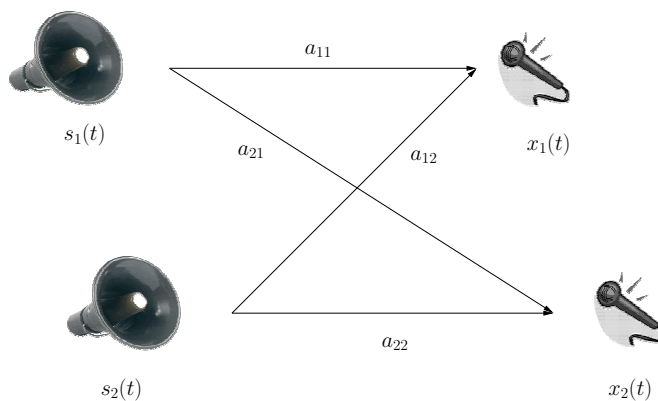


Figure 1: Cocktail party problem: case of two recorded signals generated by two source signals.

The problem can be modeled as

$$x(t, \omega) = As(t, \omega), \quad t \in \mathcal{T}, \omega \in \Omega$$

where  $A \in \mathbb{R}^{n \times n}$  is a constant unknown matrix termed *mixing matrix*,  $\mathcal{T} \subset \mathbb{R}$  is the time set,  $\Omega$  is the sample space of a probability space,  $x : \mathcal{T} \times \Omega \rightarrow \mathbb{R}^n$  is an observed stochastic process, and  $s : \mathcal{T} \times \Omega \rightarrow \mathbb{R}^n$  is a hidden stochastic process whose components  $s_1, \dots, s_n$  are statistically independent.

However, in practical applications, we only have a sampled signal  $(x(t_1), \dots, x(t_N))$  with values in  $\mathbb{R}^n$ , and we do not know the probability distribution of  $x$ , but we assume that  $x$  comes from a realization  $(s(t_1), \dots, s(t_N))$  of the random process  $s$  according to a linear law

$$x(t_i) = As(t_i), \quad i = 1, \dots, N.$$

By defining matrices

$$S := \begin{bmatrix} s_1(t_1) & s_1(t_2) & \cdots & s_1(t_N) \\ \vdots & \vdots & \ddots & \vdots \\ s_n(t_1) & s_n(t_2) & \cdots & s_n(t_N) \end{bmatrix}$$

and

$$X := \begin{bmatrix} x_1(t_1) & x_1(t_2) & \cdots & x_1(t_N) \\ \vdots & \vdots & \ddots & \vdots \\ x_n(t_1) & x_n(t_2) & \cdots & x_n(t_N) \end{bmatrix},$$

we obtain the form

$$X = AS,$$

where  $X$  is known (observed) and  $A$  and  $S$  are unknown. The goal is to recover  $A$  and  $S$  as faithfully as possible, knowing that the rows of  $S$  are realizations of statistically independent random processes.

A popular approach is to look for an *unmixing matrix*  $W \in \mathbb{R}^{n \times n}$  such that the rows of matrix

$$Z = W^T X \tag{1}$$

are “as independent as possible”. To this end, we make use of a *contrast function*  $\gamma_X : \mathbb{R}^{n \times n} \rightarrow \mathbb{R} : W \mapsto \gamma_X(W)$  which measures the “dependence” of the rows of  $Z$ . The problem then reduces to computing a matrix  $W$  that minimizes the cost function  $\gamma_X$ . This approach is termed Independent Component Analysis (ICA).

Most ICA algorithms minimize a contrast whose domain of definition is restricted to the orthogonal group  $O_n$  (instead of  $\mathbb{R}^{n \times n}$ ). This restriction is made possible by the whitening technique, as follows. Let us suppose that the rows of  $X$  are linearly independent (a reasonable assumption). Let  $W = U\Sigma V$  be a singular value decomposition of  $W$ . Equation (1) becomes

$$Z = V^T \Sigma U^T X.$$

Whitening consists in choosing matrices  $\Sigma$  and  $U$  such that the correlation matrix of  $\Sigma U^T x$  is a multiple of the identity, i.e.,  $(\Sigma U^T X)(\Sigma U^T X)^T = \sigma I$ . It then remains to choose  $V \in O_n$  that minimizes  $\gamma_{\tilde{X}}(V)$ , where  $\tilde{X} := \Sigma U^T X$ . From now on, we assume that  $X$  has been whitened (i.e.,  $XX^T = \sigma I$ ) and that it remains to find  $V \in O_n$  that minimizes the interdependence between the rows of  $Z = V^T X$ .

We propose to use a contrast function of “SOBI” type [BAMCM97] which measures the “diagonality” of lagged covariance matrices of  $z$ . Concretely, for a lag  $d$ , we define the (symmetric) covariance matrix  $R(d) \in \mathbb{R}^{n \times n}$  whose element  $(i, j)$  is given by

$$(R_X(d))(i, j) = \sum_{k=1}^{N-d} (X(i, k)X(j, k+d) + X(i, k+d)X(j, k)).$$

Since  $Z = V^T X$ , we get

$$R_Z(d) = V^T R_X(d) V. \tag{2}$$

We choose a collection of lags  $d_1, \dots, d_K$ ; from matrix  $X$  (known), we build the covariance matrices

$$C_k = R_X(d_k), \quad k = 1, \dots, K,$$

and we consider the contrast

$$f : O_n \rightarrow \mathbb{R} : V \mapsto f(V) = \sum_{k=1}^K \|\text{off}(V^T C_k V)\|_F^2 \quad (3)$$

where  $\text{off}(M)$  stands for  $M$  with the diagonal elements set to zero, and  $\|\cdot\|_F$  denotes the Frobenius norm. This contrast can be thought of as a “measure of diagonality” of matrices  $R_Z(d_k)$ ,  $k = 1, \dots, K$ , and thus gives an account of the “level of dependence” of the rows of  $Z$ .

## 2 Questions

1. Show that it is possible to choose matrices  $\Sigma$  diagonal and  $U$  orthogonal such that  $(\Sigma U^T X)(\Sigma U^T X)^T$  is a multiple of the identity.
2. Show that matrices  $R_X(d)$  are symmetric. Are they always positive definite? Prove equation (2).  
Suggestion: Write the expression of  $R_X(d)$  in matrix form.
3. Prove that the orthogonal group  $O_n = \{V \in \mathbb{R}^{n \times n} : V^T V = I\}$  is a submanifold of  $\mathbb{R}^{n \times n}$ .
4. Give an expression for  $T_V O_n$ .
5. Give an expression for the Riemannian metric induced on  $O_n$  by the canonical metric  $\langle Z_1, Z_2 \rangle = \text{trace}(Z_1^T Z_2)$  of  $\mathbb{R}^{n \times n}$ .
6. Give an expression for  $\text{grad } f(V)$ , where  $f$  is the cost function on  $O_n$  defined in (3).  
Remarks:  $\|M\|_F^2 = \text{trace}(M^T M)$ . For all matrices  $A$  and  $B$  of compatible dimensions,  $\text{trace}(\text{off}(A)\text{off}(B)) = \text{trace}(\text{off}(A)B) = \text{trace}(A\text{off}(B))$ ;  $\text{trace}(AB) = \text{trace}(BA)$ ;  $\text{trace}(A^T) = \text{trace}(A)$ .
7. Write an expression for  $\text{Hess } f(V)[Z]$  where  $Z \in T_V O_n$ .
8. Write a steepest-descent method for  $f$  and use it in the provided Matlab template.
9. Likewise with a conjugate gradient method.

## 3 Complementary questions

The following, more difficult questions involve quotient manifolds.

Let  $\mathbb{R}_*^{n \times p}$ ,  $p \leq n$ , denote the set of all full rank matrices of size  $n \times p$ . Consider the cost function

$$\bar{f} : \mathbb{R}_*^{n \times p} \rightarrow \mathbb{R} : Y \mapsto \sum_{k=1}^K (\log \det \text{ddiag}(Y^T C_k Y) - \log \det(Y^T C_k Y)), \quad (4)$$

where  $C_1, \dots, C_K$  are given symmetric positive-definite matrices.

The Hadamard inequality guarantees that  $\det(M) \leq \prod_i m_{ii}$  for all symmetric positive-definite matrices  $M$ , and that the equality holds if and only if  $M$  is diagonal. It follows that  $\bar{f}(V)$  is nonnegative for all  $V$  and is zero if and only if all the matrices  $V^T C_k V$  are diagonal.

Let  $\mathcal{D}_*^p$  be the set of nonsingular diagonal matrices. Let  $Y\mathcal{D}_*^p$  denote the set  $\{YD : D \in \mathcal{D}_*^p\}$  and let  $\mathbb{R}_*^{n \times p} / \mathcal{D}_*^p$  denote the quotient space  $\{Y\mathcal{D}_*^p : Y \in \mathbb{R}_*^{n \times p}\}$ .

1. Prove that the function  $\bar{f}$  defined in (4) satisfies the invariance property

$$\bar{f}(YD) = \bar{f}(Y)$$

for all  $D \in \mathcal{D}_*^p$ .

2. Choose a Riemannian metric on  $\mathbb{R}_*^{n \times p}$  that turns  $\mathbb{R}_*^{n \times p} / \mathcal{D}_*^p$  into a Riemannian quotient manifold.
3. Let  $f$  be the projection of  $\bar{f}$  onto the quotient  $\mathbb{R}_*^{n \times p} / \mathcal{D}_*^p$ , i.e.,  $\bar{f} = f \circ \pi$  where  $\pi$  is the natural projection associated with the quotient. Write the gradient and the Hessian of  $f$  by means of horizontal lifts.

## References

- [BAMCM97] Adel Belouchrani, Karim Abed-Meraim, Jean-François Cardoso, and Eric Moulines, *A blind source separation technique using second-order statistics*, IEEE Trans. Signal Process. **45** (1997), no. 2, 434–444.