





Probabilistic Graphical Models

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Probabilistic graphical models

- Graphical models are used in various domains:
 - Machine learning and artificial intelligence
 - Computational biology
 - Statistical signal and image processing
 - Communication and information theory
 - Statistical physics.....
- Based on correspondences between graph theory and probability theory
- Important but difficult problems:
 - Computing likelihoods, marginal distributions, modes
 - Estimating model **parameters** and **structure** from noisy data

Probabilistic Graphical Models

• Role of the graphs:

graphical representations of probability distributions

- Visualize the structure of a model
- Insights into the model properties (eg conditional independence)
- Design and motivate new models
- Design graph based algorithms for inference



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Probabilistic Graphical Models

Probability Theory

• Sum rule

$$p(\mathbf{x}) = \sum_{\mathbf{y}} p(\mathbf{x}, \mathbf{y})$$

• Product rule

$$p(\mathbf{x}, \mathbf{y}) = p(\mathbf{x}|\mathbf{y})p(\mathbf{y})$$

• From these we have Bayes' theorem

$$p(\mathbf{y}|\mathbf{x}) = \frac{p(\mathbf{x}|\mathbf{y})p(\mathbf{y})}{p(\mathbf{x})}$$

with normalization $p(\mathbf{x}) = \sum_{\mathbf{y}} p(\mathbf{x}|\mathbf{y})p(\mathbf{y})$

All probabilistic inference and learning manipulations amount to repeated application of these 2 equations

•Directed graphs: Bayesian Networks

•Conditional independence and Markov properties

•Undirected graphs: Markov Random Fields

•Inference and learning

•Some illustrations

Directed graphs

Bayesian Networks

Directed Graphs: Decomposition

• Consider an arbitrary joint distribution

• By successive application of the product rule

$$p(x, y, z) = p(x)p(y, z|x)$$

= $p(x)p(y|x)p(z|x, y)$



General Case

• Arbitrary joint distribution,

$$P(x_1,\ldots,x_n)$$

• Successive application of the product rule

 $P(x_1,...,x_n) = P(x_1)P(x_2|x_1)...P(x_n|x_1...x_{n-1})$

 Can be represented by a fully connected graph (links to all lower-numbered nodes)

Information is in the absence of links

General relationship

• Factorization property

$$P(x_1, \dots x_n) = \prod_{k=1}^n P(x_k | pa_k)$$

Where pa_k denotes the parents of x_k

• Missing link imply conditional independencies

Graph Terminology

- Directed graph G: set of nodes and directed edges
- Acyclic graph: no loop in the graph
- Parents of a node X:
 Y such that Y → X in G
- Descendent of a node X: Y that can be reached from X following directed edges



Directed Acyclic Graphs: Bayesian Networks

• The graph can be used to impose constraints on the random vector (x_1, \ldots, x_7) (ie. on the distribution P):



 $P(x_1)P(x_2)P(x_3)$ $P(x_4|x_1, x_2, x_3)$ $P(x_5|x_1, x_3)$ $P(x_{6}|x_{4})$ $P(x_7|x_4, x_5)$

No directed cycles

Bayesian Network

• A couple (p, G) so that

$$p(x_1,\ldots,x_n) = \prod_k p(x_k | pa_k^G)$$

$$p \sim \mathcal{L}(G)$$

Hidden variables

• Variables may be hidden (latent) or visible (observed)



 Latent variables may have a specific interpretation, or may be introduced to permit a richer class of distribution

Example 1: Mixtures of Gaussians

• Linear super-position of K Gaussians

$$P(y) = \sum_{k=1}^{K} \pi_k \mathcal{N}(y|\mu_k, \sigma_k^2)$$

• Normalization and positivity require

$$\sum_{k=1}^{K} \pi_k = 1 \qquad 0 \leqslant \pi_k \leqslant 1$$

• illustration: mixture of 3 Gaussians



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Latent Variable Viewpoint

- Discrete latent variable $x \in \{1, \dots K\}$ describing which component generated data point y
- Conditional distribution of observed variable

$$P(y|X=k) = \mathcal{N}(y|\mu_k, \sigma_k^2)$$

• Prior distribution of latent variable

$$P(X=k) = \pi_k$$



• Marginalizing over the latent variable we obtain

$$P(y) = \sum_{k=1}^{K} \pi_k \mathcal{N}(y|\mu_k, \sigma_k^2)$$

Example 2: State Space Models

Hidden Markov chain

$$\sum_{k=1}^{n} b(x^{k}) x^{k} (x^{k}) b(x^{k}) (x^{k}) (x^{k}) - \cdots + b(x^{k}) b(x^{k}) (x^{k}) - \cdots + b(x^{k}) b(x^{k}) b(x^{k}) (x^{k}) - \cdots + b(x^{k}) b(x^{k$$

• Kalman filter



Frequently wish to solve the problem of computing

$$P(x_t|y_1,\ldots,y_n)$$

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Causality

- Directed graphs can express causal relationships
- Often we observe child variables and wish to infer the posterior distribution of parent variables
- Example:



• Note: inferring causal structure from data is subtle

Conditional independence and Markov properties

Conditional independence

• X independent of Y given Z if for all values of z,

$$P(x|y,z) = P(x|z)$$

• Notation:

$$X\perp Y|Z$$

• Equivalently

$$P(x, y|z) = P(x|y, z)P(y|z)$$
$$= P(x|z)P(y|z)$$

• Conditional independence crucial in practical applications since we can rarely work with a general joint distribution

Difference between dependence and conditional dependence

- Traffic jams and snowmen are correlated
- But conditionally on snow falls, the size of the traffic jams and the number of snowmen are independent



The concept of conditional dependence is more suited than dependence to capture « direct » dependencies between variables

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Markov properties

- Can we determine the conditional independence properties of a distribution directly from its graph?
- YES: "d-separation", one subtleties due to the presence of head-to-head nodes, *explaining away effect*



Example 1: Tail-to-head node

• Joint distribution

$$P(a, b, c) = P(a)P(c|a)P(b|c)$$



 $a \not\perp b$ (c not observed)

$$P(a, b|c) = P(a|c)P(b|c) \implies a \perp b|c \quad (c \text{ observed})$$

• An observed c blocks the path from a to b

Example 2: Tail-to-tail node

• Joint distribution

$$P(a, b, c) = P(c)P(a|c)P(b|c)$$

 $a \not\perp b$ (c not observed)

$$P(a, b|c) = P(a|c)P(b|c) \implies a \perp b|c \quad (c \text{ observed})$$

• An observed c blocks the path from a to b

b

C

a

Example 3: "Explaining Away" (V-structure)

Illustration: pixel colour in an image



An observed I unblocks the path from S to L

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$$p(G = 1|B = 1, F = 1) = 0.8$$

$$p(G = 1|B = 1, F = 0) = 0.2$$

$$p(G = 1|B = 0, F = 1) = 0.2$$

 $p(G = 1|B = 0, F = 0) = 0.1$

$$B$$
 F G

$$egin{array}{rll} p(B=1) &=& 0.9 \ p(F=1) &=& 0.9 \end{array}$$

and hence p(F=0) = 0.1

$$p(F=0|G=0) = \frac{p(G=0|F=0)p(F=0)}{p(G=0)}$$
$$\simeq 0.257$$

The probability of an empty tank increased by observing G = 0.

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$$p(F = 0|G = 0, B = 0) = \frac{p(G = 0|B = 0, F = 0)p(F = 0)}{\sum_{F \in \{0,1\}} p(G = 0|B = 0, F)p(F)}$$

$$\simeq 0.111$$

- The probability of an empty tank is **reduced** by observing B = 0. This is referred to as "explaining away".

- B and F are **negatively correlated** conditioned on G despite being independent

d-separation: Consider 3 groups of nodes A, B, C

To determine whether $A \perp B | C$ is true, consider all possible paths from any node in A to any node in B

- $A \perp B | C$ true if all paths from A to B are blocked by C
- Any such path is blocked if there is a node X which is head-to-tail or tail-to-tail with respect to the path and X is in C
 Or

if the node is **head-to-head** and neither the **node nor any** of **its descendants** is in C



Undirected graphs

Markov Random Fields

Undirected graphical models

- The second major class of graphical models
- Graphs specify **factorizations** of distributions and sets of conditional independence relations (**Markov properties**)
- Markov Random Fields or Markov network

Cliques and maximal cliques

• A clique C is a subset of vertices all joined by edges



- Cliques: (1), (2),(12), (23).....
- Maximal cliques: (123), (345), (456), (47)

$$p(w, x, y, z) = \frac{1}{Z} \psi_A(w, x, y) \psi_B(x, y, z)$$

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Undirected Graphs: Factorization

 Provided p(x) > 0 then joint distribution is product of non-negative functions over the *cliques* of the graph

$$p(\mathbf{x}) = \frac{1}{Z} \prod_{C} \psi_{C}(\mathbf{x}_{C})$$

 $X = \{X_i, i \in V\} \qquad X_C = \{X_i, i \in C\}$

• where $\psi_C(\mathbf{x}_C)$ are the *clique potentials*, and *Z* is a normalization constant

Undirected graphs: conditional independencies

Conditional independence given by graph separation
 x independent of y given z



Conditional independencies: Markov properties

Terminology: Markov blanket or Markov Boundary of a node \mathcal{X}_i is the set of nodes N(i) such that

$$P(x_i|x_{-i}) = P(x_i|x_{N(i)})$$

or equivalently $X_i \perp X_{-i \cup N(i)} | X_{N(i)}$



Markov blankets on the graph

- Directed case: Parents, Children, Co-parents
- Undirected case: Neighbors





Markov property: for X (p) wrt G

- Graph G=(V,E)
- $X = \{X_i, i \in V\}$ random vector
- $X_A = \{X_i, i \in A\}$
- X is Markov wrt G

if X_A and X_B are conditionally independent given X_C whenever C separates A and B

 Specifying conditional independencies using the neighborhood N(i) is enough (V finite)
Undirected graphs: Markov Networks

The law of the random variable $X = (X_1 \dots X_n)$ is a graphical model according to the non-directed graph G

if for all i :

$$X_i \perp \{X_j, j \notin N(i) \cup \{i\}\} \mid \{X_j, j \in N(i)\}$$

We can write $\mathcal{L}(X) \sim G$

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Connection with directed acyclic graphs

The Moral graph gets the parents married

The moral graph Gm associated to a directed acyclic graph G is obtained by:

- Setting an edge between each parent of each nodes
- Replacing arrows by edges

We have:

$$\mathcal{L}(X) \sim \mathbf{G} \implies \mathcal{L}(X) \sim \mathbf{Gm}$$

Hammersley-Clifford theorem

In practice (**computation**), we use the connection between conditional independencies (Markov properties) and **factorization** property

• Boltzmann-Gibbs representation

$$\Psi_c(x_c) = \exp(-E(x_c))$$

• P is a positive MRF (satisfies Markov properties) is equivalent to P is a Gibbs distribution

$$P(x) = \frac{1}{Z} \exp(-E(x))$$

• Energy function

$$E(x) = \sum_{c} E_c(x_c)$$

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Example: pairwise Markov Random Fields

• Cliques: pairs, singletons

$$E(x) = \sum_{i} \{ \Psi_i(x_i) + \frac{1}{2} \sum_{j \in N(i)} \Psi_{ij}(x_i, x_j) \}$$

• Famous ones:

Ising model: binary variables on a graph G with pairwise interactions

$$P(x;\theta) = \frac{1}{Z} \exp\left(\sum_{i} \theta_{i} x_{i} + \sum_{i \sim j} \theta_{ij} x_{i} x_{j}\right)$$

- Potts model: K-ary variables

Interaction parameters+ external field parameters

Example: graph representation of a Pairwise MRF

• Typical application: image region labelling



Illustration: image segmentation

site/vertex i: pixel, y_i : observed grey level, x_i : label/0 or 1/ binary variable





Probabilistic Graphical Models

Challenging computational problems

- Frequently, it is of interest to compute various quantities associated with an undirected graphical model:
 - The log normalization constant log Z
 - Local marginal distributions (p(xi)) or other local statistics
 - Modes and most probable configurations
- Often grow rapidly with graph size and max clique size
- Example: Computing the normalization constant for binary random variables

$$Z = \sum_{x \in \{0,1\}^n} \prod_{c \in C} \psi_c(x_c)$$

Complexity scales exponentially as 2^n

Probabilistic Graphical Models

Inference and learning

Inference in Graphical models

- Exploit the graphical structure to find efficient algorithm for inference and to make the structure of these algorithms clear (eg propagation of local messages around the graph)
- Exact inference
- Approximate inference

Inference

• Simple example: Bayes' theorem



Message Passing: compute marginals



• Find marginal for a particular node

$$p(x_i) = \sum_{x_1} \dots \sum_{x_{i-1}} \sum_{x_{i+1}} \dots \sum_{x_L} p(x_1, \dots, x_L)$$

- for *M*-state nodes, cost is $O(M^L)$
- exponential in length of chain
- but, we can exploit the graphical structure (conditional independencies)

Message Passing

• Joint distribution

$$p(x_1,\ldots,x_L)=\frac{1}{Z}\psi(x_1,x_2)\ldots\psi(x_{L-1},x_L)$$

• Exchange sums and products: ab+ ac = a(b+c)

$$m_{\alpha}(x_{i}) \quad \text{before } x_{i}$$

$$p(x_{i}) = \frac{1}{Z} \cdots \sum_{x_{2}} \psi(x_{2}, x_{3}) \left[\sum_{x_{1}} \psi(x_{1}, x_{2}) \right]$$

$$\cdots \sum_{x_{L-1}} \psi(x_{L-2}, x_{L-1}) \left[\sum_{x_{L}} \psi(x_{L-1}, x_{L}) \right]$$

$$m_{\beta}(x_{i}) \quad \text{after } x_{i}$$

Message Passing

• Express as product of messages

• Recursive evaluation of messages: Linear in L

$$m_{\alpha}(x_{i}) = \sum_{x_{i-1}} \psi(x_{i-1}, x_{i}) m_{\alpha}(x_{i-1})$$
$$m_{\beta}(x_{i}) = \sum_{x_{i+1}} \psi(x_{i}, x_{i+1}) m_{\beta}(x_{i+1})$$

• Find Z by normalizing $p(x_i)$

Belief Propagation

- Extension to general tree-structured graphs
- At each node:
 - form product of *incoming* messages and local evidence
 - marginalize to give *outgoing* message
 - one message in each direction across every link



• Fails if there are loops

Junction Tree Algorithm

- An efficient exact algorithm for a general graph
 - applies to both directed and undirected graphs
 - compile original graph into a tree of cliques
 - then perform message passing on this tree
- Problem:
 - cost is exponential in size of largest clique
 - many vision models have intractably large cliques

Loopy Belief Propagation

- Apply belief propagation directly to general graph
 - possible because message passing rules are local
 - need to keep iterating
 - might not converge
- State-of-the-art performance in some applications

Max-product Algorithm: most probable x

Goal: find

$$\mathbf{x}^{\mathsf{MAP}} = \arg\max_{\mathbf{x}} p(\mathbf{x})$$

$$\phi(x_i) = \max_{x_1} \cdots \max_{x_{i-1}} \max_{x_{i+1}} \cdots \max_{x_L} p(x_1, \dots, x_L)$$

- then

$$x_i^{\mathsf{MAP}} = \arg\max_{x_i} \phi(x_i)$$

- Message passing algorithm with "sum" replaced by "max"
- Example:
 - Viterbi algorithm for HMMs

Inference and learning

In general: Hidden or latent X (underlying scene) and Observed Y (image)

- Inference: computing P(x|y) ("posterior")
- Learning: computing P(y) (likelihood) usually
 (*θ* : parameter estimation based on ML)

 $P_{\theta}(y)$

Likelihood of the data y

$$L(\theta) = P_{\theta}(y)$$

Maximum (log) likelihood

$$\theta_{ML} = \arg \max_{\theta} \log L(\theta)$$

Example: classification with context

• The labeling problem

★ n objects/individuals
$$(i \in V = \{1, ..., n\})$$

★ K labels $(k \in \mathcal{A} = \{1, ..., K\})$
★ n * ... observations $(y = (y_1, y_2, ...))$

assign a label to each object consistently with y: $\mathbf{x}: V \to \mathcal{A}$

$$x = (x_1, \ldots, x_n \in \mathcal{A}^n)$$

(assignment, colouring (graph), configuration (random fields)

Contextual constraints: distance, similarity, compatibility, etc.

- Image analysis, segmentation, etc.
- Biometrics: spatially related observations
- Documents analysis: hyperlinks between documents



No context

Too much context

Good compromise

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Connection Cost/Energy and probability

★ assignment cost x : V → A
c(i, k) [likelihood of k at site i] or c_y(i, k) [data term]
★ Neighborhood cost:
i and j nearby ⇒ x_i and x_j similar/compatible
→ graph G = (V, E): if (i, j) ∈ E
→ cost w_{ij} × d_{ij}(x_i, x_j) [Ψ_{ij}(x_i, x_j)]

Total cost:
$$E(x) = \sum_{i \in S} c(i, x_i) + \sum_{(i,j) \in E} w_{ij} d_{ij}(x_i, x_j)$$

- Goal: find x that maximizes E
- Discrete optimization, NP-hard, find approximations, satisfying assignments

Optimal configuration for Pairwise MRF with energy E

Energy and MAP rule

• Corresponding graphical model: Pairwise MRF

$$E(x) = \sum_{i} \{ \Psi_{i}(x_{i}) + \frac{1}{2} \sum_{j \in N(i)} \Psi_{ij}(x_{i}, x_{j}) \}$$

• Maximum A Posteriori (MAP) principle:

$$\hat{x} = \arg \max_{x \in \mathcal{A}^n} P(x|y)$$

Hidden MRF: accounting for observations

- Observations, eg. Measures $Y = \{Y_i, i \in S\}$
- Hidden data, eg. Labels, X discrete MRF $P(x) = \frac{1}{Z} \exp(-E(x))$
- Data term, $P(y|x) = \exp(-E(y|x))$

Conditional MRF (posterior): $P(x|y) = \frac{1}{Z_y} \exp(-E_y(x))$

$$E_y(x) = E(x) + E(y|x)$$

E(x): Regularizing term (prior, context)E(y | x): Data term

MAP so

Nution
$$\hat{x} = \arg\min_{x \in \mathcal{L}^n} E_y(x)$$

Probabilistic Graphical Models

Approximate solutions

- Deterministic approaches: relaxation, variational methods (mean field, etc.)
- Stochastic approaches: Gibbs sampling, simulation methods (MC)
- Classification approaches: hard clustering, ICM, K-means
- Parameter estimation approaches: soft clustering, EM

Approximate Inference

For general graphical models (not tree-structured)

$$p(x) = \frac{1}{Z} \prod_{c \in \mathcal{C}} \exp(\Psi_c(x_c))$$

All basic computations are intractable, combinatorial for large G

Likelihood and partition function $Z = \sum_{x \in \mathcal{X}^N} \prod_{c \in \mathcal{C}} \exp(\Psi_c(x_c))$ Marginals and conditionals $p(x_j) = \frac{1}{Z} \sum_{x_i, i \neq j} \prod_{c \in \mathcal{C}} \exp(\Psi_c(x_c))$ Modes $x^* = \arg \max_{x \in \mathcal{X}^N} \prod_{c \in \mathcal{C}} \exp(\Psi_c(x_c))$

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Approximate Inference

- Stochastic (Sampling)
 - Metropolis-Hastings, Gibbs, (Markov Chain) Monte Carlo, etc
 - Computationally expensive, but is "exact" (in the limit)
- Deterministic (Optimization)
 - Mean Field (MF), Loopy Belief Propagation (LBP)
 - Variational Bayes (VB), Expectation Propagation (EP)
 - Computationally cheaper, but is not exact (gives bounds)

True distribution

Monte Carlo

VB / Loopy BP / EP

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Taxonomy of inference methods



General View of Variational Inference

- Consider arbitrary distribution q(x) over the latent variables
- The following decomposition always holds

$$\log p(y|\theta) = F(q,\theta) + KL(q,p)$$

where

$$F(q,\theta) = \sum_{x} q(x) \log \frac{p(x,y|\theta)}{q(x)}$$
$$KL(q,p) = -\sum_{x} q(x) \log \frac{p(x|y,\theta)}{q(x)}$$

Decomposition

Maximizing over q(x) would give the true posterior distribution – but this is intractable by definition



Factorized Approximation

- Goal: choose a family of distributions which are:
 - sufficiently flexible to give good posterior approximation
 - sufficiently simple to remain tractable
- Here we consider factorized distributions

$$q(x) = \prod q_i(x_i)$$

- No further assumptions ⁱare required!
- Optimal solution for one factor, keeping the remained fixed

$$q_j^*(x_j) \propto \exp(I\!\!E_{q_{\backslash j}^*}[\log p(y, x)]) \qquad q_{\backslash j}^* = \prod_{i \neq j} q_i^*$$

• Coupled solutions so initialize then cyclically update

Factorized approximation

In practice, we compute $q_j^*(x_j) \propto \exp(I\!\!E_{q^*_{\backslash j}}[\log p(x|y)])$ ommiting terms that does not depend on \mathcal{X}

and hope to recognize a standard distribution or normalize

Ex. Hidden Markov Field

$$p(y|x) = \prod_{i} p(y_{i}|x_{i})$$

$$p(x) \text{ is a MRF so that } p(x_{j}|x_{\setminus j}) = p(x_{j}|x_{N(j)})$$

$$\implies p(x|y) \propto p(y|x) \ p(x) \propto \prod_{i} p(y_{i}|x_{i}) \ p(x_{j}|x_{\setminus j})p(x_{\setminus j})$$

$$\implies q_{j}^{*}(x_{j}) \propto \exp(I\!\!E_{q_{\setminus j}^{*}}[\log p(y_{j}|x_{j}) + \log p(x_{j}|X_{\setminus j})]))$$
apprinting terms that does not depend on T i

ommiting terms that does not depend on ${\mathcal X}_j$

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Example: Discrete Hidden MRF

$$p(x) = \frac{1}{Z} \exp(E(x)) \text{ with } x_i \in \{1 \dots K\} \text{ and } E(x) = \sum_{i \sim j} \Psi_{ij}(x_i, x_j)$$
$$\Rightarrow p(x_j | x_{\backslash j}) \propto \exp(\sum_{i \in N(j)} \Psi_{ij}(x_i, x_j))$$

and

$$p(y|x) = \prod_{i} p(y_i|x_i) \text{ with } p(y_i|x_i = k) = f_{\theta_k}(y_i)$$

$$q_j^*(x_j) \propto p(y_j|x_j) \exp(\mathbb{E}_{q_{N(j)}^*}[\sum_{i \in N(j)} \Psi_{ij}(x_i, x_j)])$$

$$q_j^*(x_j) \propto p(y_j|x_j) \exp(\sum_{i \in N(j)} \mathbb{E}_{q_i^*}[\Psi_{ij}(x_i, x_j)])$$

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Illustration: Ising model, binary MRF

$$\begin{split} \Psi(x_i, x_j) &= \theta_{ij} \ x_i x_j \qquad x_i \in \{-1, 1\} \\ \text{Remark: } \Psi(x_i, x_j) &= \theta_{ij} (2x_i - 1)(2x_j - 1) \text{ if } x_i \in \{0, 1\} \end{split}$$

$$I\!\!E_{q_i^*}[\Psi(x_i, x_j)] = \theta_{ij} \; x_j I\!\!E_{q_i^*}[x_i] = \theta_{ij} \; x_j \; (q_i^*(x_i = 1) - q_i^*(x_i = -1))$$

$$q_j^*(x_j = 1) = \frac{1}{1 + \frac{p(y_j | x_j = -1)}{p(y_j | x_j = 1)}} \exp\left(-2\sum_{i \in N(j)} \theta_{ij} \left(q_i^*(x_i = 1) - q_i^*(x_i = -1)\right)\right)}$$
$$q_j^*(x_j = -1) = 1 - q_j^*(x_j = 1)$$

Fixed point equation or iterative updating

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Iterated Conditional Modes (ICM) for HMRF

[Besag 70s]

For each j in turn

$$x_{j}^{*} = \arg\max_{x} p(y_{j}|x_{j} = x) p(x_{j} = x|x_{N(j)}^{*})$$
$$x_{j}^{*} = \arg\max_{x} p(y_{j}|x_{j} = x) \exp(x \sum_{i \in N(j)} \theta_{ij} x_{i}^{*})$$

A modal version of variational mean field

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Gibbs sampler for HMRF

[Geman & Geman 80s]

A stochastic version of ICM or a simulated version of variational Mean Field

For each j in turn $x_j^* \sim p(y_j | x_j) \ p(x_j | x_{N(j)}^*)$

$$x_j^* \sim p(y_j|x_j) \exp(x_j \sum_{i \in N(j)} \theta_{ij} x_i^*)$$
 (Ising)

Sample
$$u \sim Uniform(0,1)$$

 $x_{j}^{*} = 1$ if $u \leq \frac{1}{1 + \frac{p(y_{j}|x_{j}=-1)}{p(y_{j}|x_{j}=1)}} \exp(-2\sum_{i \in N(j)} \theta_{ij} x_{i}^{*})$

 $x_i^* = -1$ otherwise

Probabilistic Graphical Models

Sampling vs Variational approximations

True distribution

Monte Carlo

VB / Loopy BP / EP

- 1) MCMC (eg Gibbs sampler)
- Theoretical properties
- High computational cost
- Complicated convergence monitoring
- Model selection & general noise model: not straightforward

2) Variational (eg VEM)

- Fast and flexible
- Lack of theoretical properties
- Global covariance structure cannot be estimated

Example 1: MRI Brain scan segmentation

Assign each voxel to a class (label) (among K classes) [Forbes et al 2011]



Graphical model representation



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Cooperative segmentation of tissues and structures

observations





No anatomical information



Meilleure segmentation incontestable des putamens et des thalamus

Cooperative method

Example 2: texture recognition

• Learning step: model estimation







• Interest points



neighborhood graph



• Test step: classification









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Example 2: texture recognition



Thank you for your attention

