



# Inverse problems in functional brain imaging

## Joint detection-estimation in fMRI

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1: CEA/NeuroSpin/LNAO

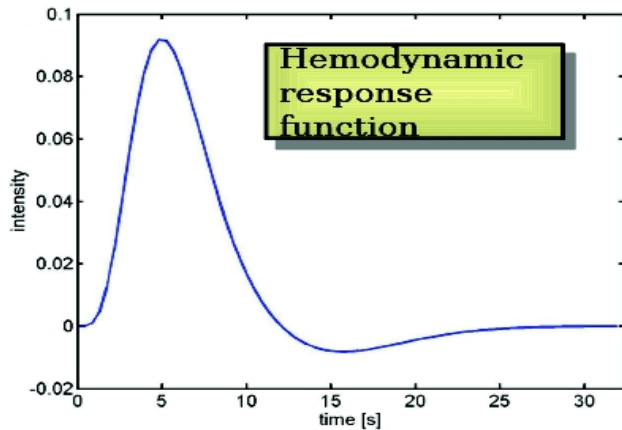
2: IFR49



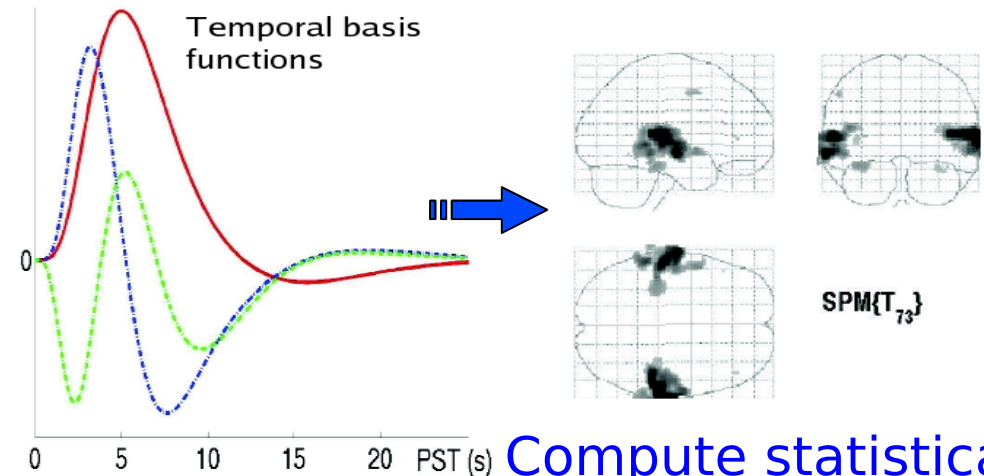
# Classical fMRI analysis

## 1. **Detect** and **localise** brain activations

Ex: In SPM [Friston et al, 1994], the BOLD response is modelled with:



or

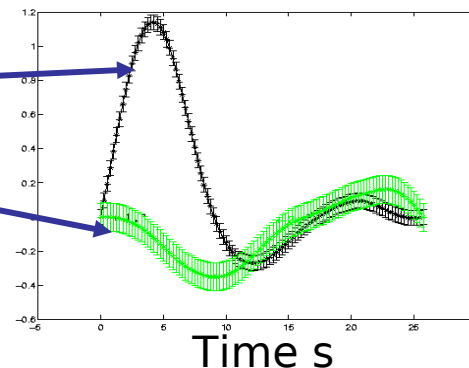
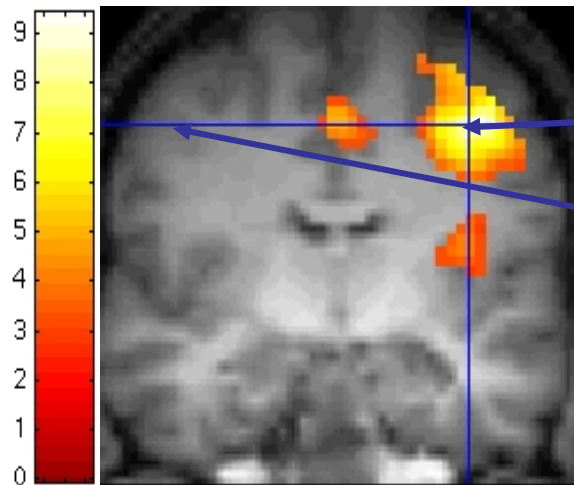


## 2. **Estimate** the dynamics of activation

[Goutte et al, IEEE TMI 2000; Marrelec et al, HBM 2003]

Compute statistical activation Maps

HRF estimates



Probe brain dynamics



# Classical fMRI analysis

- GLM limitations:
  - A single HRF shape is not physiologically appropriate
  - Studies report HRF variability:
    - within subject (between regions, sessions, conditions, trials)
      - [Miezin et al., NIM 2000; Ciuciu et al., IEEE TMI 2003]
      - [Neumann et al, NIM 2003, 2006; Smith et al, NIM 2005]
    - between subjects
      - [Handwerker et al., NIM 2004; Aguirre et al., NIM 1998]
    - between groups (infants, patients,...)
      - [D'Esposito et al, NIM 1998, 2003; Richter and Richter, NIM 2003]
    - ...

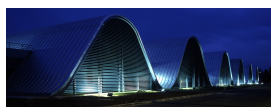


# Motivations

- **Detection** of brain activation and **estimation** of brain dynamics are addressed separately
- **Any detection** method supposes a given HRF shape
- **Any estimation** algorithm provides relevant results in activated voxels or regions only

➔ Address these two problems simultaneously in a joint detection-estimation (JDE) framework

[Makni et al., IEEE SP 2005, ISBI'06, ICASSP'06, NIM, 2008; Vincent et al, ICASSP'07, EMBC'07, ISBI'08; Ciuciu et al, ISBI'08 ; Risser et al, sub. to NIPS'08; Vincent et al, sub. to IEEE TMI]

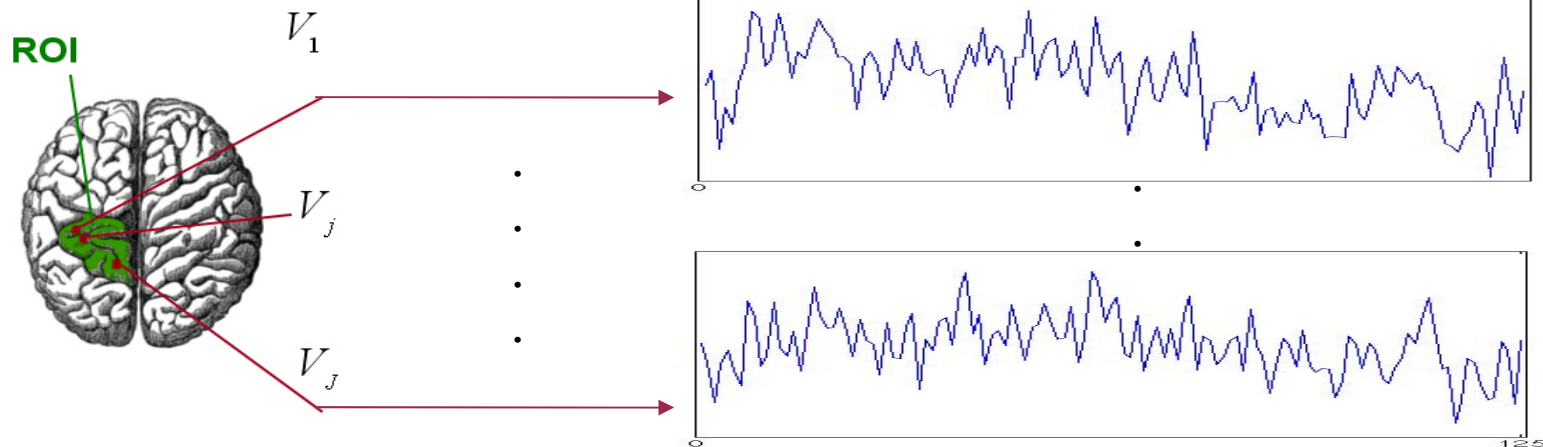


# Forward BOLD signal model

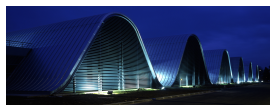
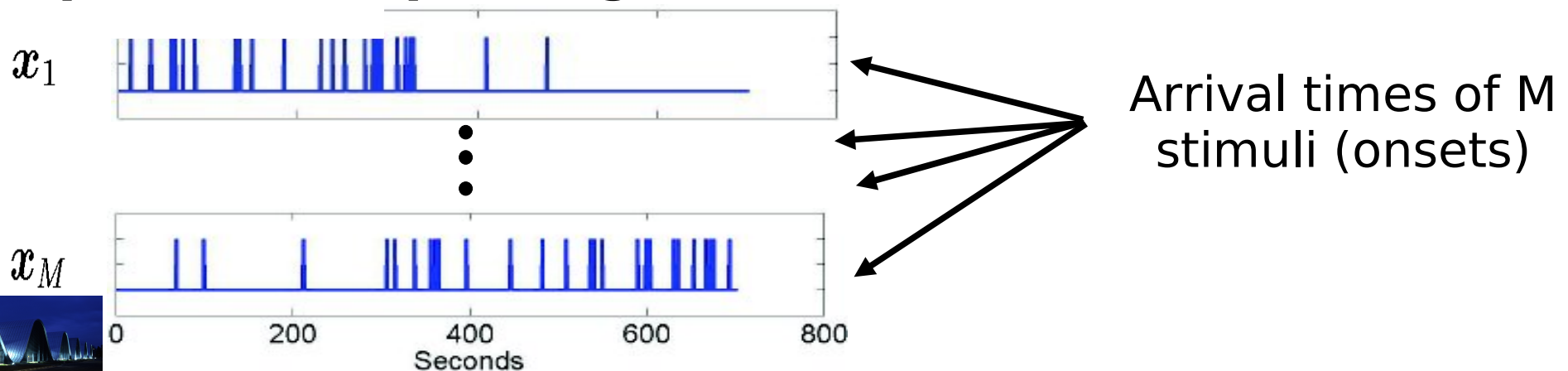
[Makni, Ciuciu et al., IEEE SP 2005;  
Makni et al, Ciuciu, NeuroImage, 2008]

- **Region definition:**

Region of interest  $\mathcal{R} = (V_j)_{j=1:J}$  ← voxels

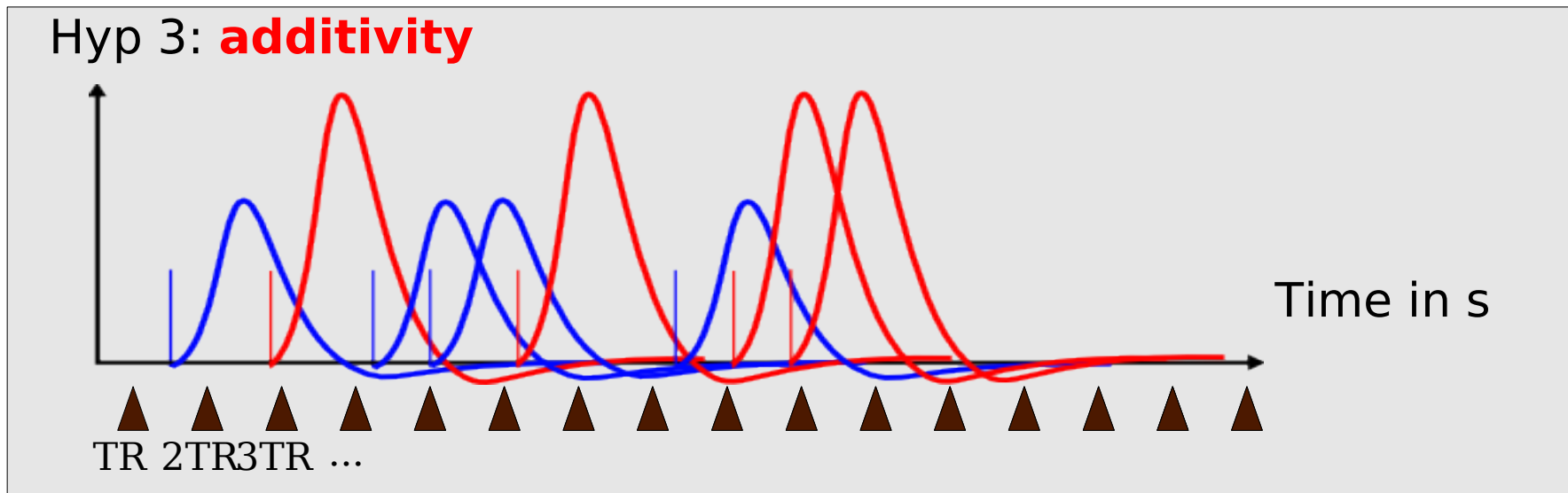
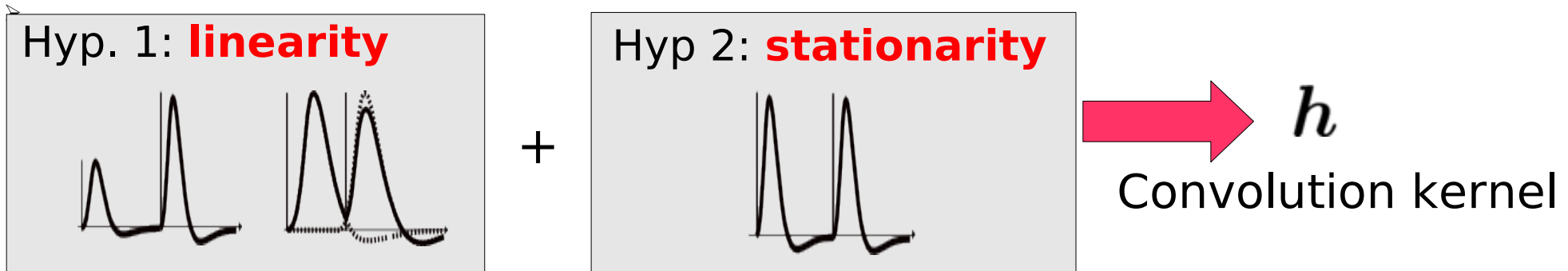


- **Experimental paradigm:**

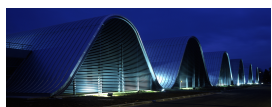


# Forward BOLD signal model

- **Temporal hypotheses:** for standard ISIs



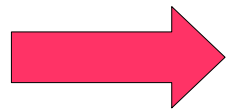
→ Stimulus-varying NRLs:  $a^1 \neq a^2$





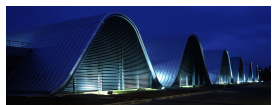
# Forward BOLD signal model

- **Spatial hypotheses:** functionally homogeneous ROI
  - › Single HRF shape
  - › Voxel-dependent magnitudes of the BOLD response



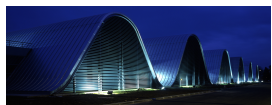
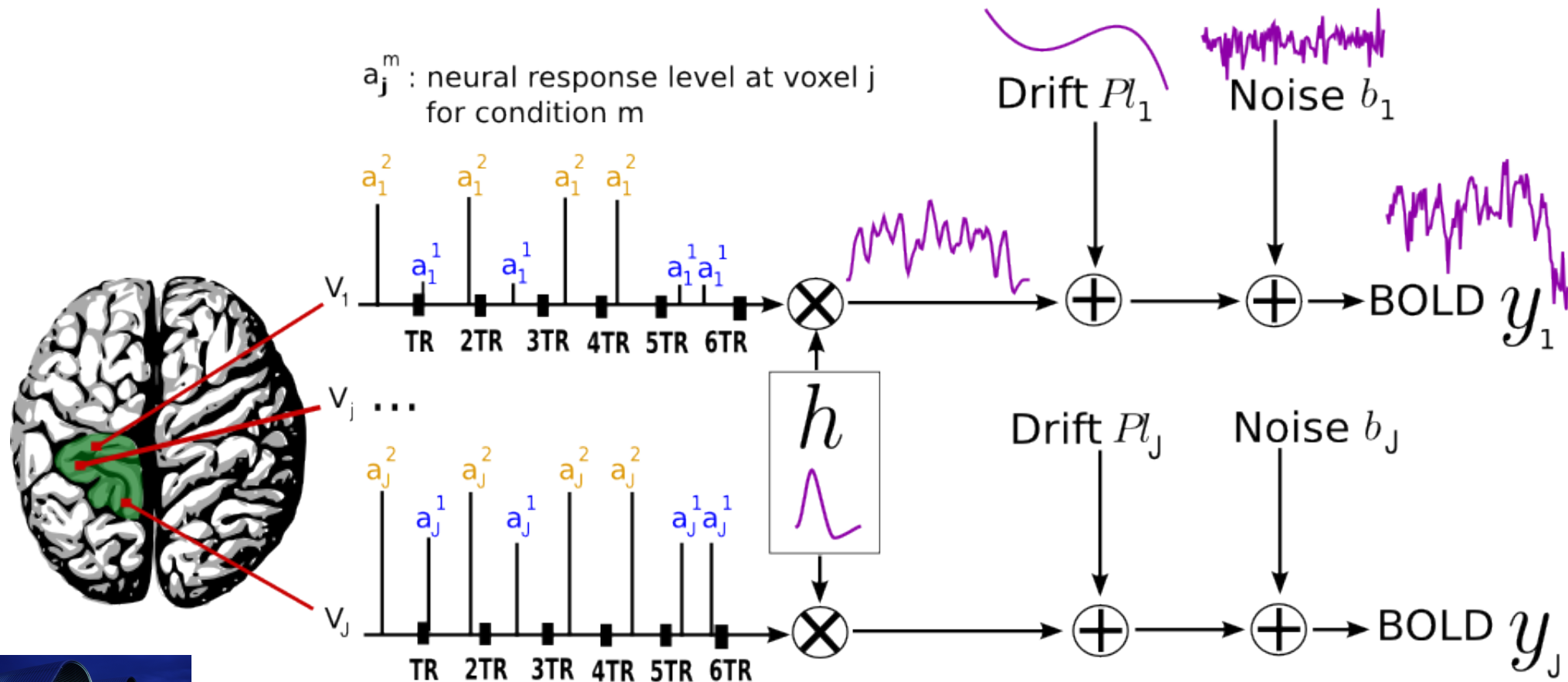
**Neural Response Levels**

$$\mathbf{a} = \{a_j^m\}_{j=1:J, m=1:M}$$



# Forward BOLD signal model

- **Spatial hypotheses:** functionally homogeneous ROI
  - Single HRF shape [Makni, Ciuciu et al., IEEE SP 2005; Makni et al, Ciuciu, NeuroImage, 2008]
  - Voxel-dependent magnitudes of the BOLD response





# Forward BOLD signal model

## Unknown parameters

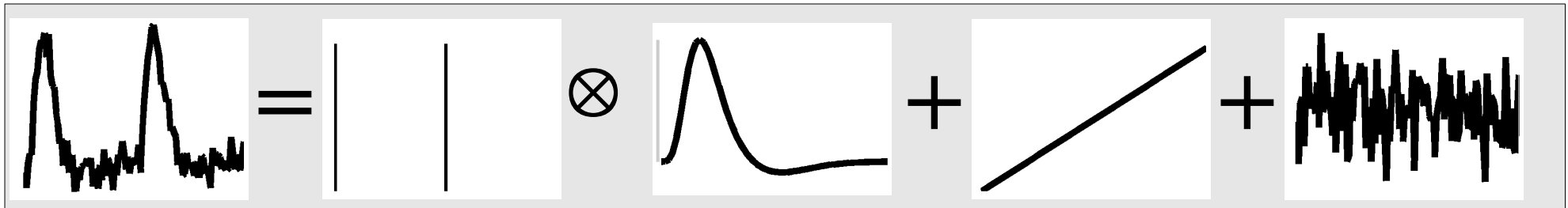
NRL in  $V_j$  and for condition  $m$

HRF

Drift coefficients

Noise statistics in voxel  $V_j$

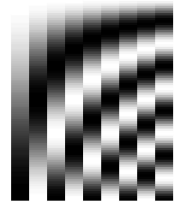
$$\mathbf{y}_j = \sum_{m=1}^M a_j^m \mathbf{X}^m \mathbf{h} + \mathbf{P} \boldsymbol{\ell}_j + \mathbf{b}_j$$



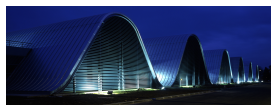
BOLD signal measured in voxel  $V_j$

Arrival time of stimulus  $m$

Orthonormal basis for low frequency drift modelling



Known parameters



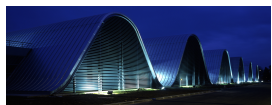
# Bayes' rule

$$p(\mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta} | \mathbf{y}) = \frac{\overset{\text{likelihood}}{p(\mathbf{y} | \mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta})} p(\mathbf{h} | \sigma_h^2) p(\mathbb{1} | \sigma_l^2) p(\mathbf{a} | \boldsymbol{\theta}) p(\boldsymbol{\theta})}{p(\mathbf{y})}$$



How the data are generated from the parameters?

**Forward modeling**



# Likelihood definition

- **Main hypothesis: noise decorrelated in space**

→ fMRI time series are statistically independent in space:

$$p(\mathbf{y} | \mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta}_0) = \prod_j p(\mathbf{y}_j | \mathbf{h}, \mathbf{a}_j, \ell_j, \theta_{0,j})$$

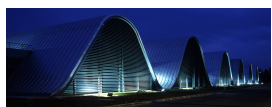
$$\propto \prod_j f_{B_j}(\mathbf{y}_j - \sum_{m=1}^M a_j^m \mathbf{X}_m \mathbf{h} - \mathbf{P} \ell_j)$$

Temporal noise model: either **white** or **serially correlated AR(1)**


$$\mathbf{b}_j \sim \mathcal{N}(\mathbf{0}, \epsilon_j^2 \mathbf{I}) \quad \rightarrow \quad \theta_{0,j} = [\epsilon_j^2]$$

$$\mathbf{b}_j \sim \mathcal{N}(\mathbf{0}, \epsilon_j^2 \boldsymbol{\Lambda}_j^{-1}) \quad \rightarrow \quad \theta_{0,j} = [\epsilon_j^2, \rho_j]$$

[Makni et al, Ciuciu, NeuroImage 2008]



# Bayes' rule

$$p(\mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta} | \mathbf{y}) = \frac{p(\mathbf{y} | \mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta}) \overbrace{p(\mathbf{h} | \sigma_h^2) p(\mathbb{1} | \sigma_l^2) p(\mathbf{a} | \boldsymbol{\theta}) p(\boldsymbol{\theta})}^{\text{Prior}}}{p(\mathbf{y})}$$


What do we know about the parameters before the data are acquired?

## Prior modeling

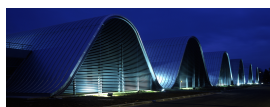
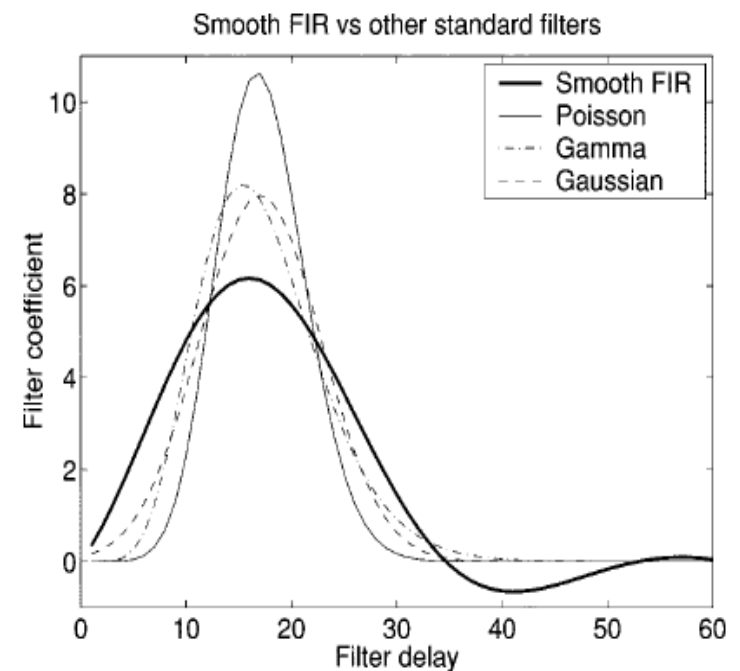


# HRF prior modeling

- **Information about the HRF shape**
  - nonparametric approaches
  - Averaging: **[Buckner et al, 1996]**
  - Selective averaging: **[Dale et al, HBM 1997]**
  - FIR (Finite Impulse Response)
  - Regularized FIR (smoothing prior):

$$h \sim \mathcal{N}(\mathbf{0}, \sigma_h^2 \Sigma)$$

**[Marrelec, Ciuciu et al, IPMI'03  
Ciuciu et al., IEEE TMI 2003]**



# NRL prior modeling

- **Spatial information** [Vincent, Ciuciu et al, ICASSP'07]

- Independence between conditions:  $p(\mathbf{a}) = \prod_m p(\mathbf{a}^m | \boldsymbol{\theta}^m)$
- **Spatial mixture model** (SMM) for each  $m$

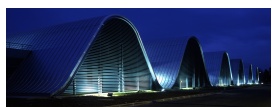
$$p(\mathbf{a}^m | \boldsymbol{\theta}^m) = \sum_{\mathbf{q}^m = \mathbf{i}} \left( \prod_{j=1}^J f_i(a_j^m | \boldsymbol{\theta}^m) \right) \Pr(\mathbf{q}^m | \beta^m)$$

$$f_i(\cdot | \boldsymbol{\theta}^m) = f(\cdot | q_j^m = i, \boldsymbol{\theta}^m), \quad \forall j$$

- Discrete Markov random Field (eg Potts/Ising):

$$\Pr(\mathbf{q}^m | \beta^m) = \frac{1}{Z(\beta^m)} \exp\left(\beta^m \sum_{j \sim k} I(q_j^m = q_k^m)\right)$$

**Partition function:**  $Z(\beta^m) = \sum_{\mathbf{q}^m} \exp\left(\beta^m \sum_{j \sim k} I(q_j^m = q_k^m)\right)$



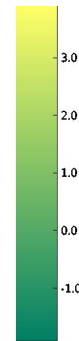
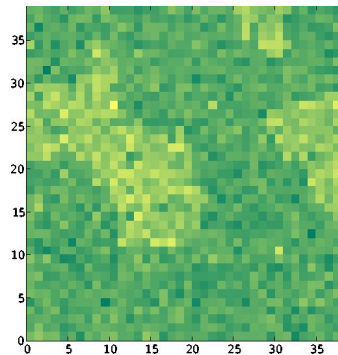
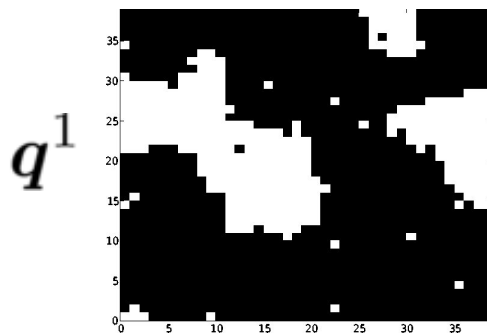
# Prior model

- **Two-class** Gaussian mixture [Vincent et al., ICASSP' 07]

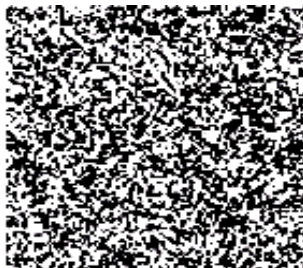
Non-activating voxels:  $f_0(\cdot | \theta^m) = \mathcal{N}(0, \sigma_{0,m}^2)$

Activating voxels:  $f_1(\cdot | \theta^m) = \mathcal{N}(\mu_{1,m}, \sigma_{1,m}^2)$

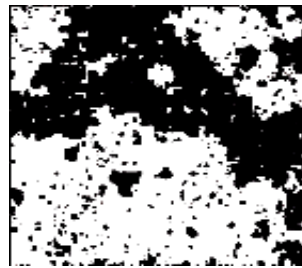
Unknown hyper-parameters:  $\{\beta^m, \sigma_{0,m}^2, \mu_{1,m}, \sigma_{1,m}^2\}$



- Influence of  $\beta$



$\beta = 0.01$



$\beta = 0.45$



$\beta = 0.91$



# Other statistical parameters

- **Noise** parameters [Makni et al, ISBI'06]

$$p(\boldsymbol{\theta}_0) = \prod_{j=1}^J p(\rho_j, \sigma_{\varepsilon_j}^2) = \prod_{j=1}^J \epsilon_j^{-1} \mathbb{1}_{(-1,1)}(\rho_j),$$

- **Mixture** probabilities [Makni et al, NIM 2008]

- 2-class mixture (Jeffreys prior):  $p(\boldsymbol{\lambda}) = \prod p(\lambda_m) \propto \prod_m \lambda_{1,m}^{1/2} \lambda_{0,m}^{1/2}$
- 3-class mixture:  $p(\boldsymbol{\lambda}) = \prod_m \mathcal{D}_3(\boldsymbol{\lambda}_m | \boldsymbol{\delta}), \quad \boldsymbol{\delta} = \delta \mathbf{1}_3$

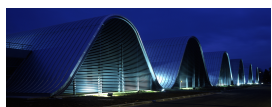
- **Mixture** parameters [Makni et al, NIM 2008]

Non-activating voxels

Activating voxels

$$p(\boldsymbol{\sigma}_{0,m}) = \prod_m \sigma_{0,m}^{-1}$$

$$p(\boldsymbol{\mu}_{1,m}, \boldsymbol{\sigma}_{1,m}) = \prod_m \mathcal{N}(\boldsymbol{\mu}_{1,m}; 0, c) \mathcal{IG}(\sigma_{1,m}^2; a, b)$$



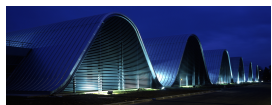


# Bayes' rule

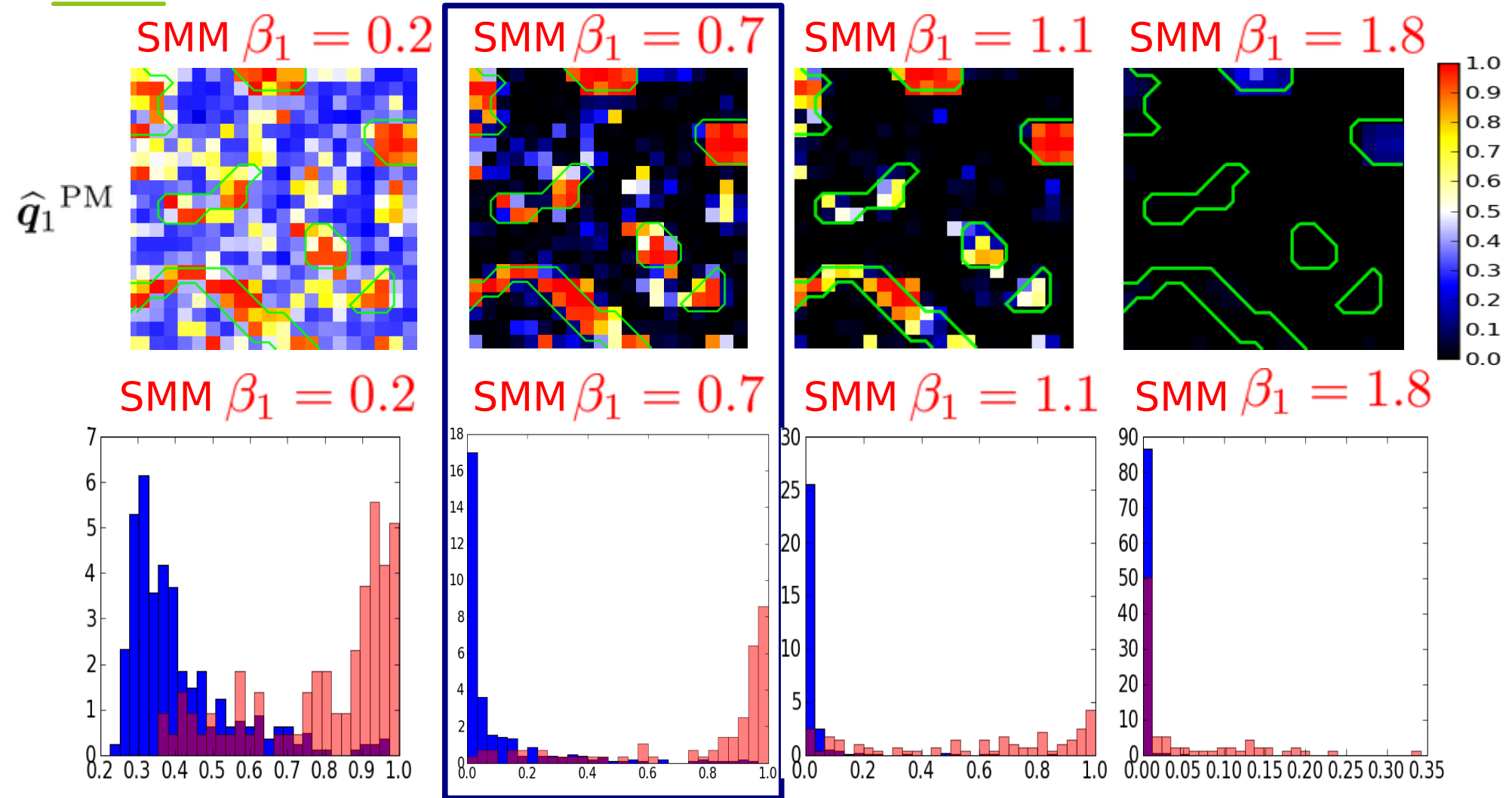
$$p(\mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta} | \mathbf{y}) \propto p(\mathbf{y} | \mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta}) p(\mathbf{h} | \sigma_h^2) p(\mathbb{1} | \sigma_l^2) p(\mathbf{a} | \boldsymbol{\theta}) p(\boldsymbol{\theta})$$

What do we know about the HRF, the NRLs and the hyper-parameters given the data?

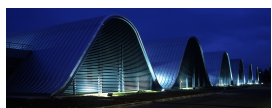
**Keystone of learning scheme:  
simulating realizations of  $p(\mathbf{h}, \mathbf{a}, \boldsymbol{\theta} | \mathbf{y})$  using Gibbs  
sampler**



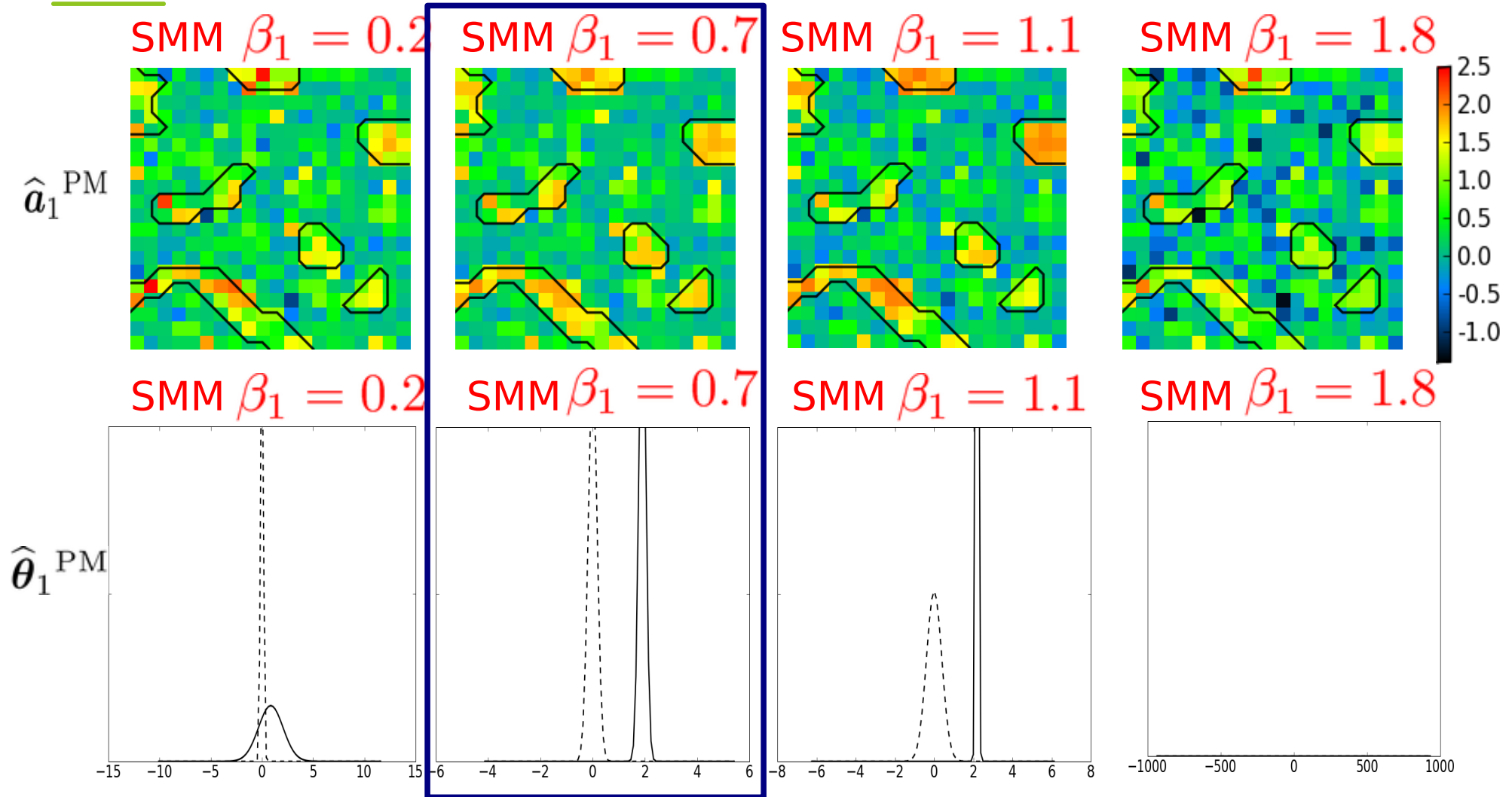
# Simulation results



Strong influence of regularization parameter



# Simulation results



dash/continuous line: density of inactivating/activating voxels



# Simulation results

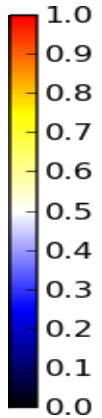
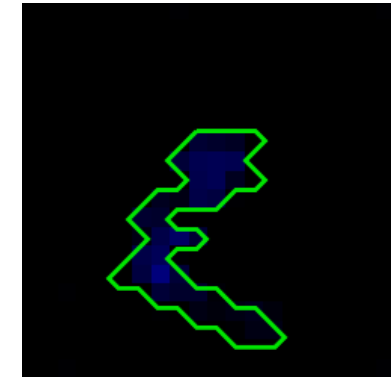
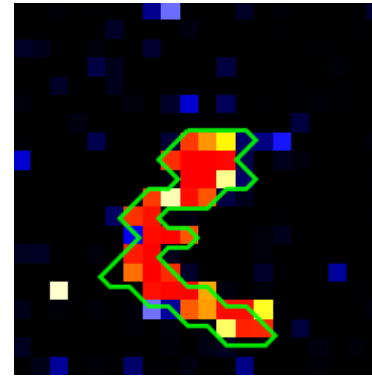
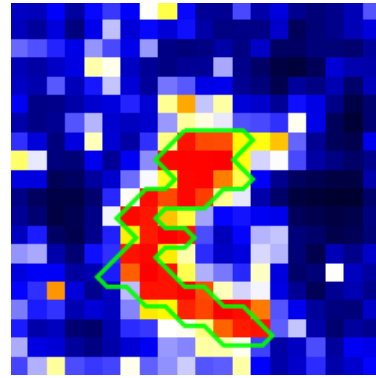
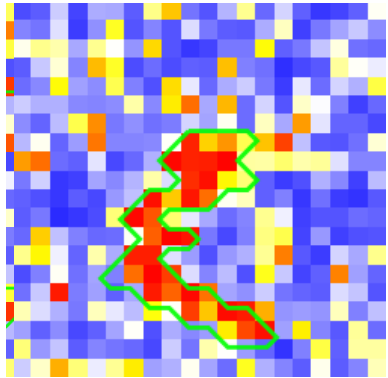
SMM  $\beta_2 = 0.2$

SMM  $\beta_2 = 0.7$

SMM  $\beta_2 = 1.1$

SMM  $\beta_2 = 1.8$

$\hat{q}_2^{\text{PM}}$

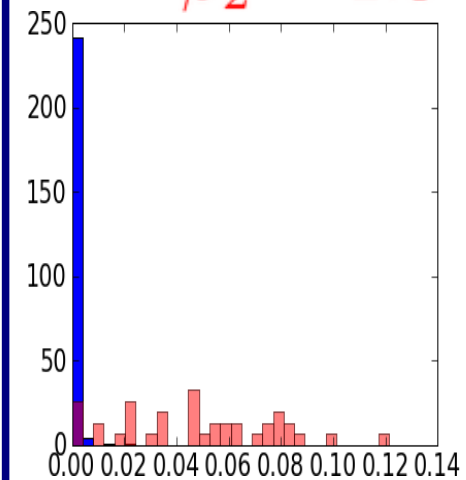
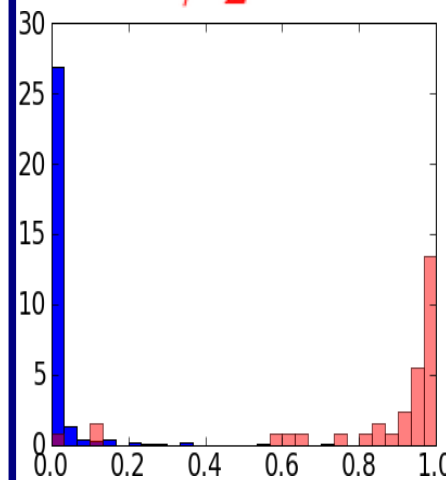
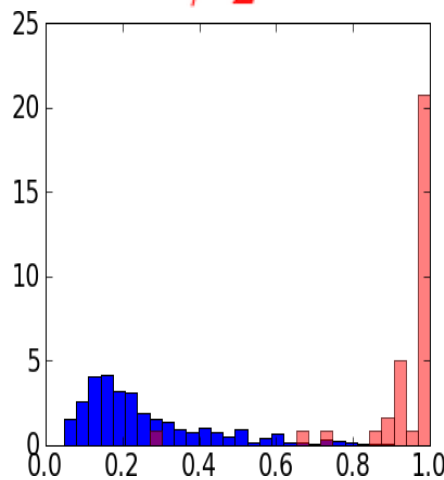
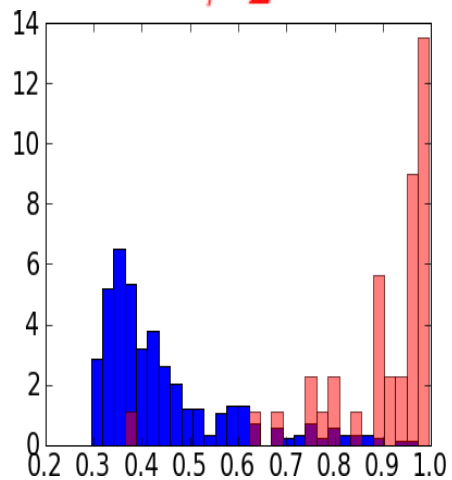


SMM  $\beta_2 = 0.2$

SMM  $\beta_2 = 0.7$

SMM  $\beta_2 = 1.1$

SMM  $\beta_2 = 1.8$



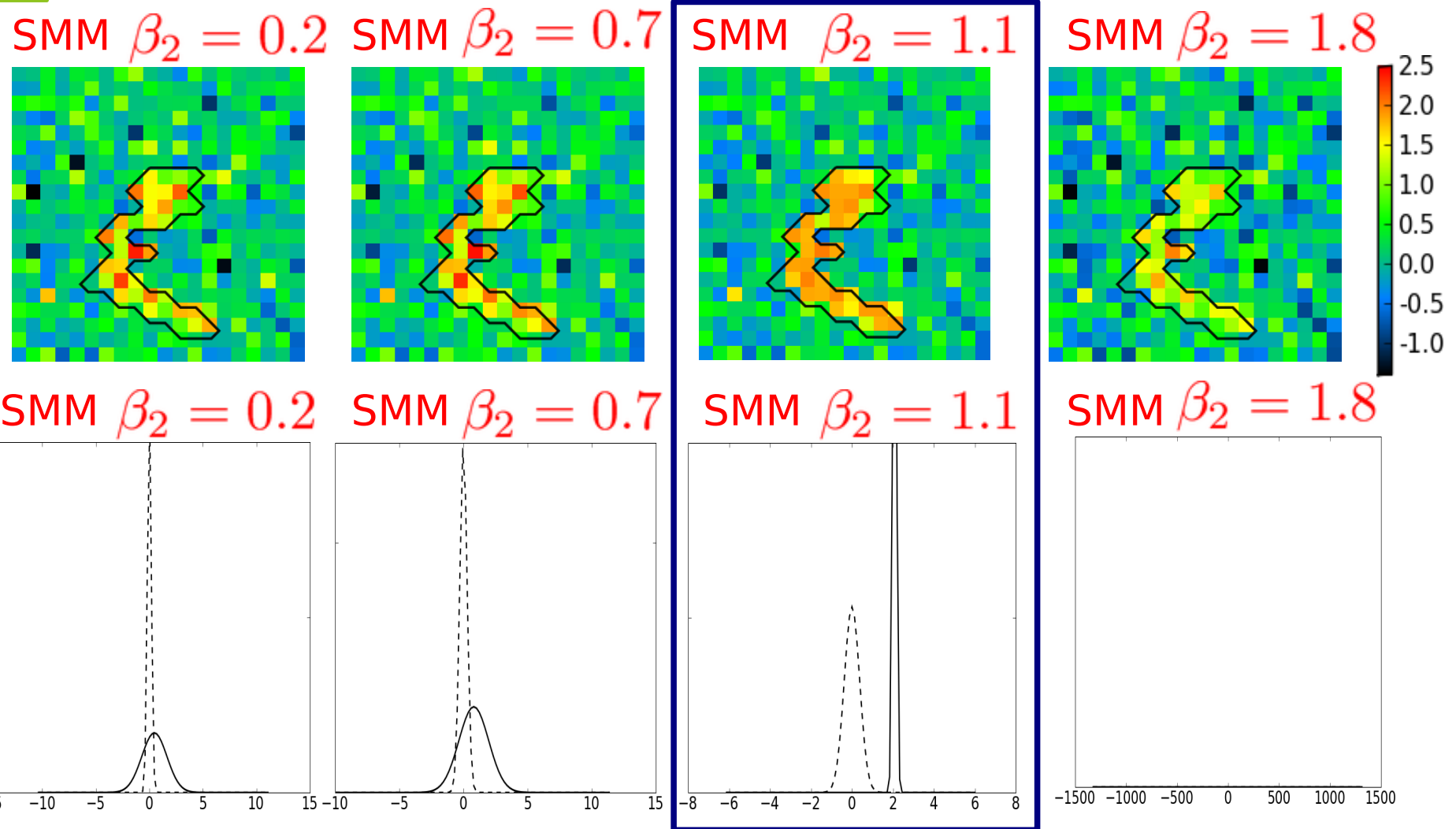
True activations

True inactivations

Best tuning: larger  $\beta_2$   $\rightarrow$  lower SNR for  $m = 2$



# Simulation results

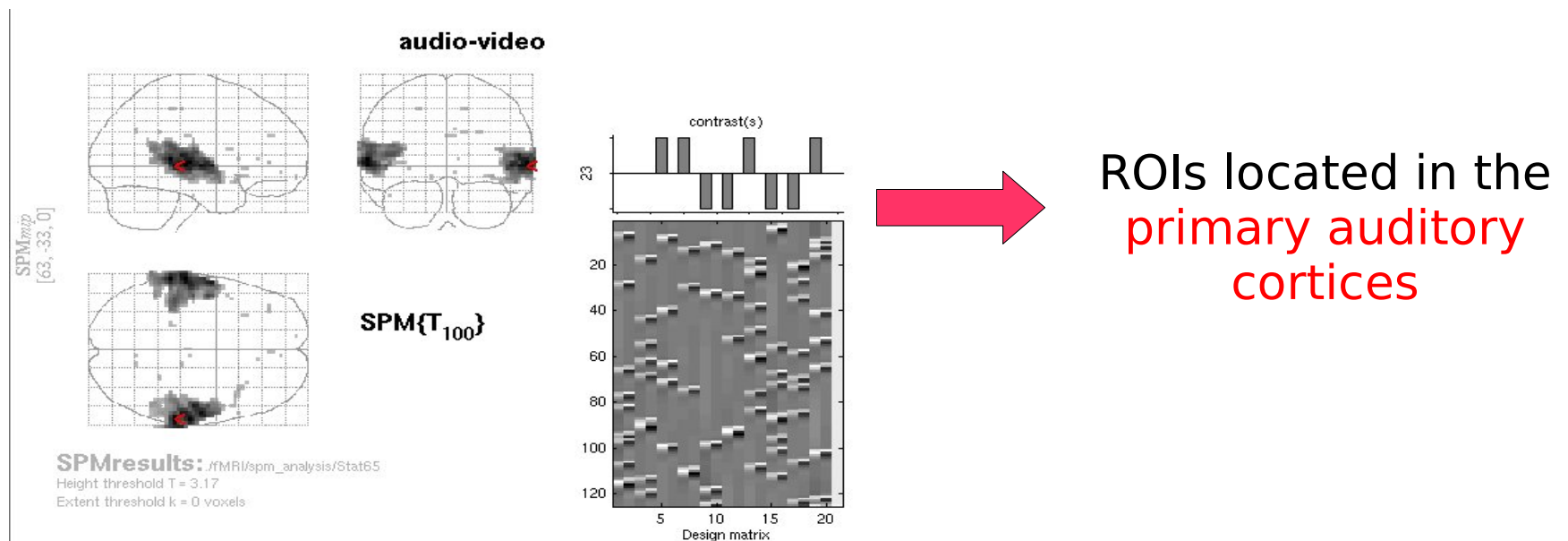


dash/continuous line: density of inactivating/activating voxels



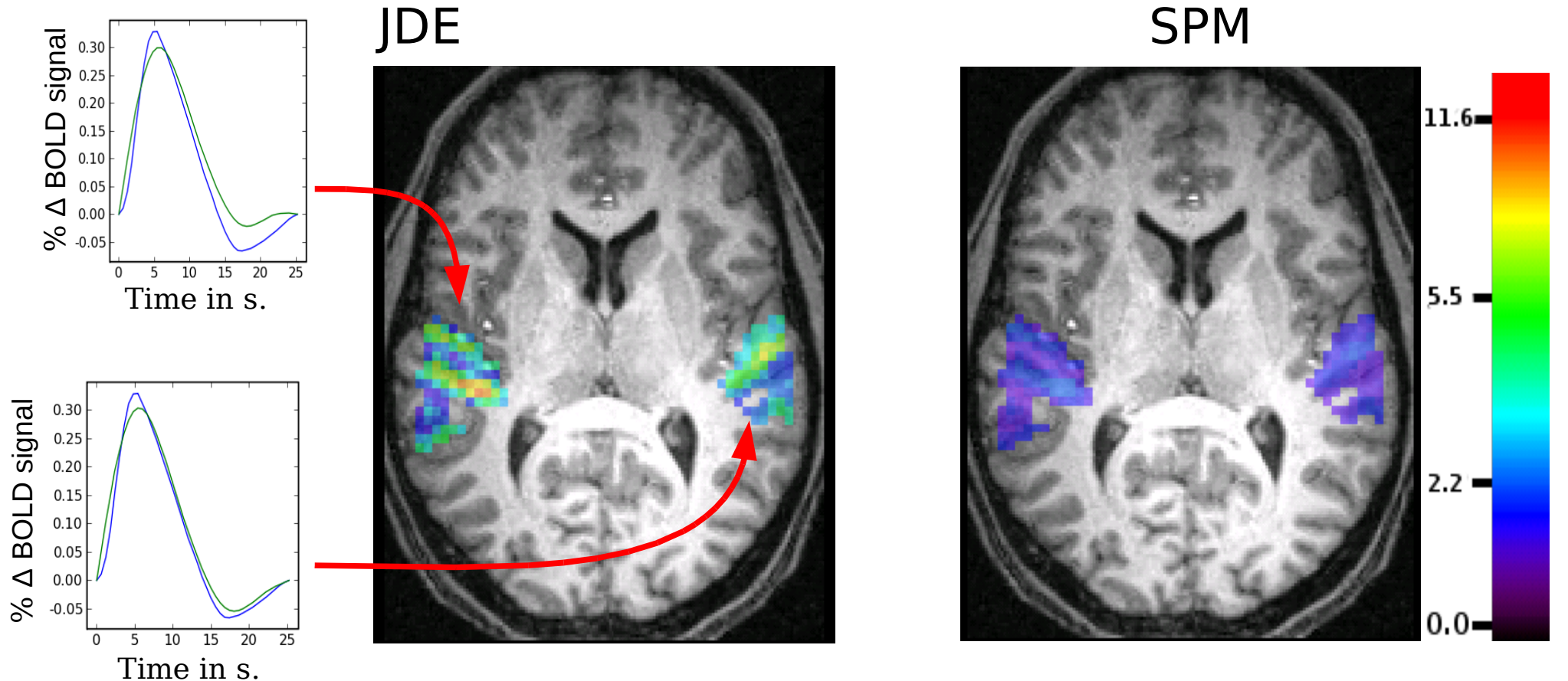
# Real data sets

- Localizer fMRI experiment:
  - Experimental conditions under study: auditory and visual stimuli
  - Event-related paradigm :
    - Short stimuli duration
    - Inter-stimulus interval :  $\sim 3s$  to  $10s$
    - Randomised sequence
  - 125 scans with  $TR = 2.4s$ , scanning at 3T

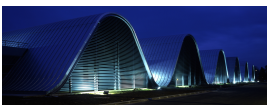


# SPM vs. JDE

Auditory - Visual contrast:  $\hat{a}^1 - \hat{a}^2$



Bilateral activation detected along the gray matter from **raw** data sets (spatially unsmoothed)



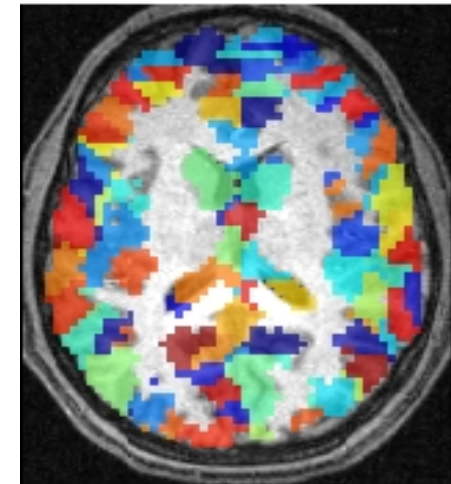
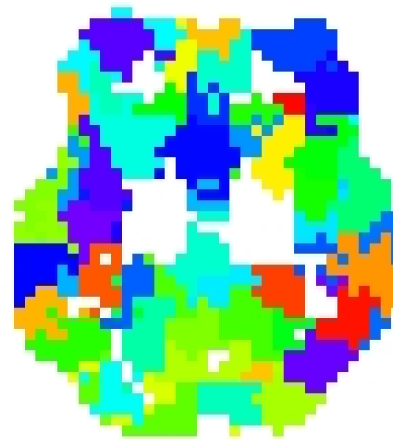
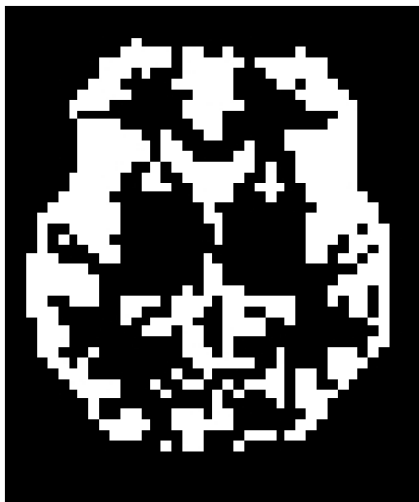
# Whole brain analysis

- **First step:**

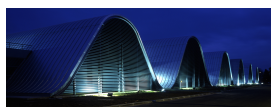
Segmentation of the gray/white matter interface from the T1 MRI

- **Second step :**

brain parcellation based upon functional similarities and spatial connectivity [Flandin et al, ISBI'02; Thirion et al, HBM 2006]



[Makni et al, Ciuciu, NIM 2008]

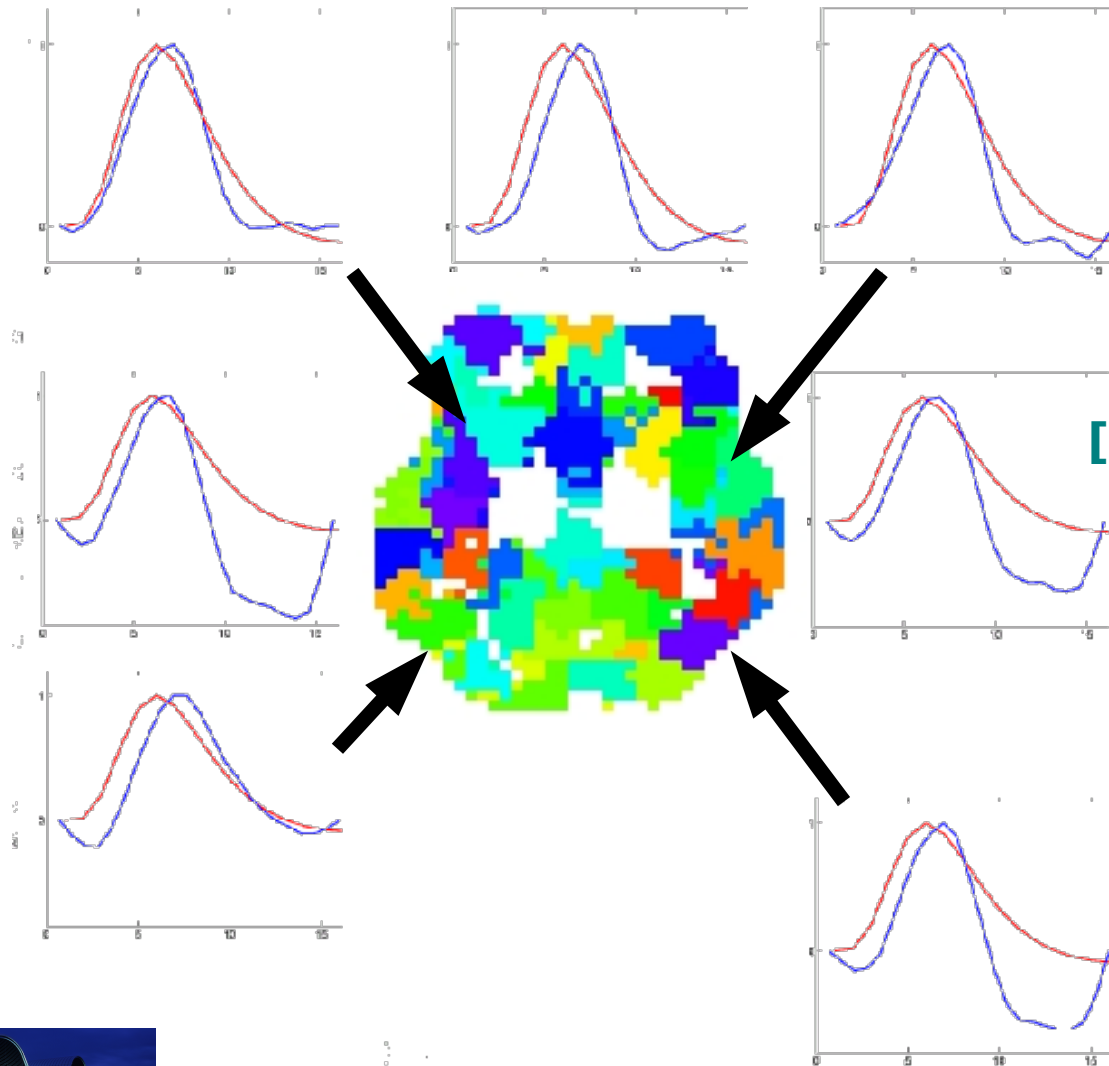




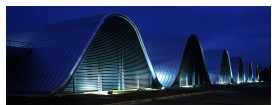
# Sources of variability

**Subject, region**

[Birn et al, NIM 2001; Marrelec et al, HBM 2003; Menon et al NIM 2003; Neumann et al NIM 2003; Handwerker et al, NIM 2004]



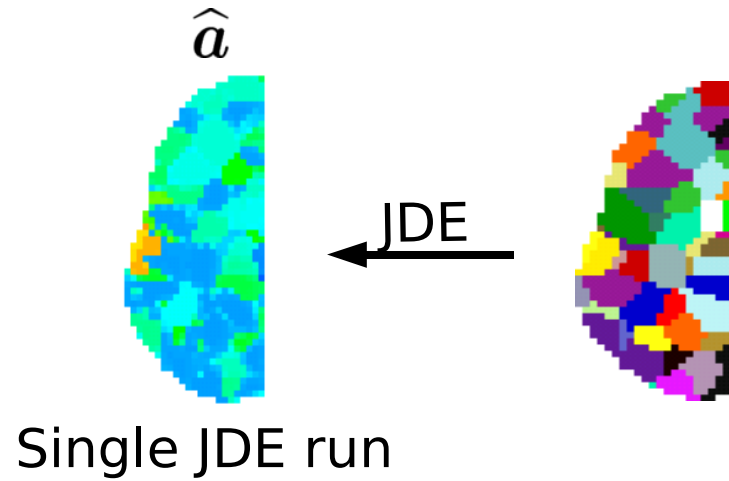
[Makni et al, Ciuciu, NIM 2008]



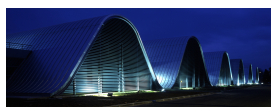
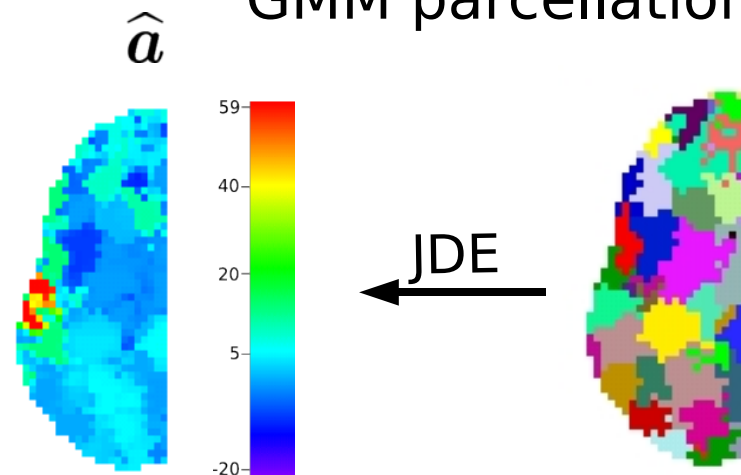
# Sensitivity analysis

[Vincent, Ciuciu, ISBI 2008]

Random parcellation



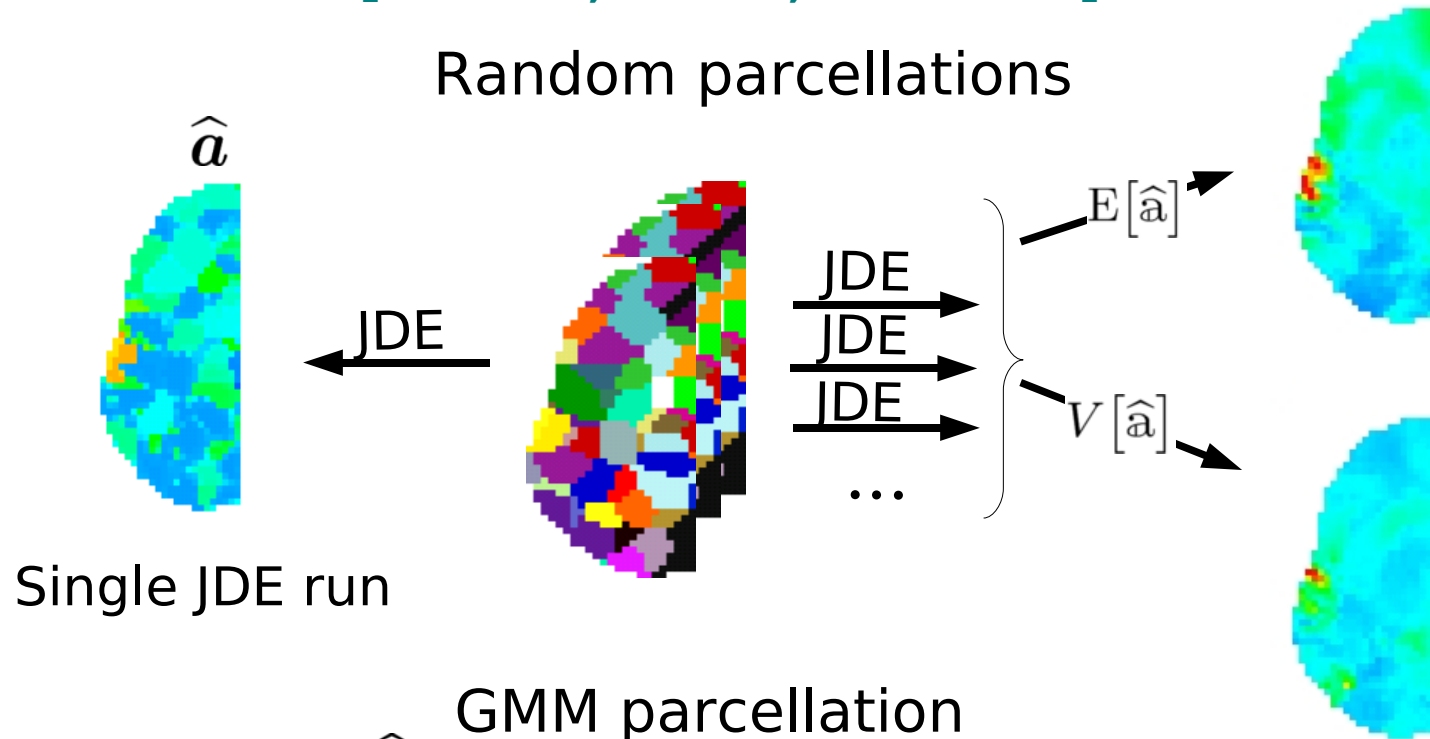
GMM parcellation



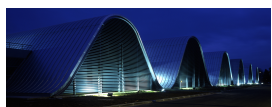
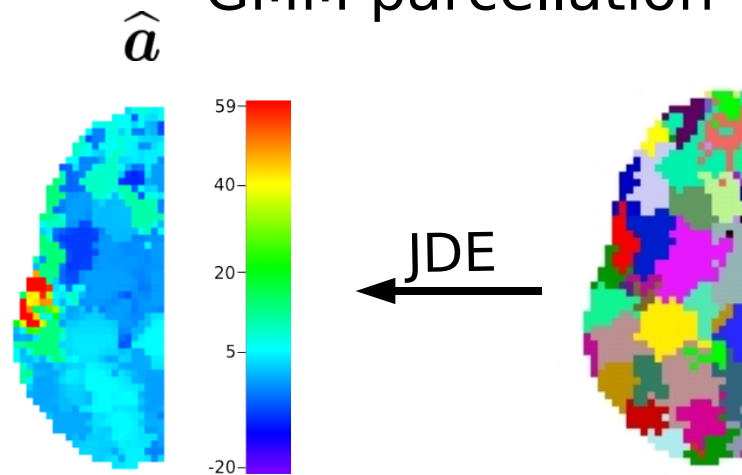
# Sensitivity analysis

[Vincent, Ciuciu, ISBI 2008]

Random parcellations

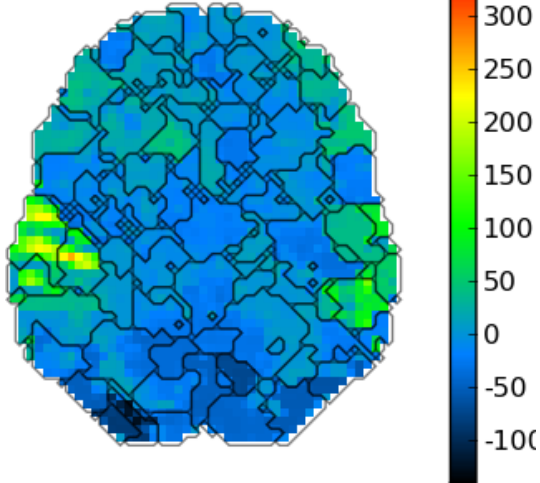


GMM parcellation

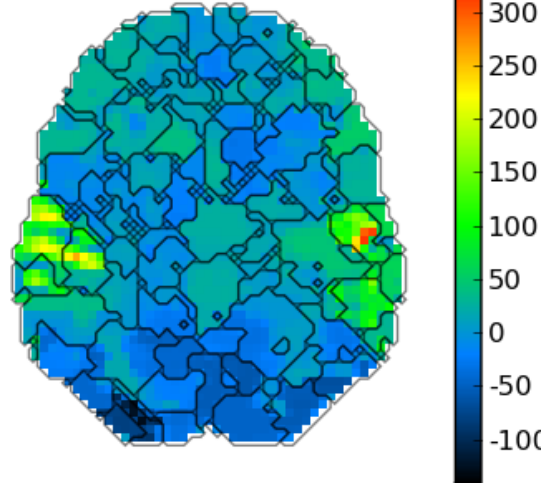


# Whole brain: supervised SMM

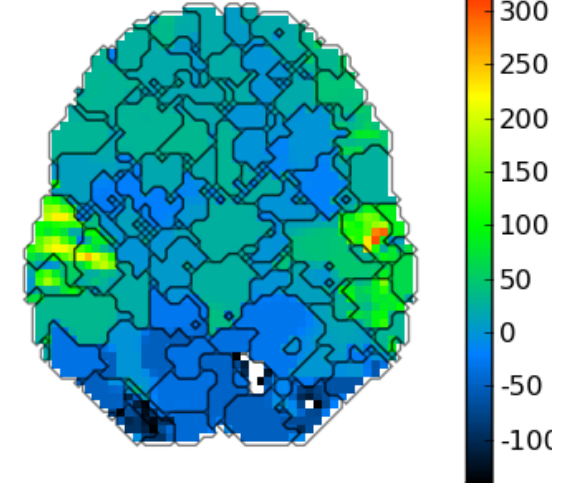
$$\hat{a}^1 - \hat{a}^2 \quad \beta=0.1$$



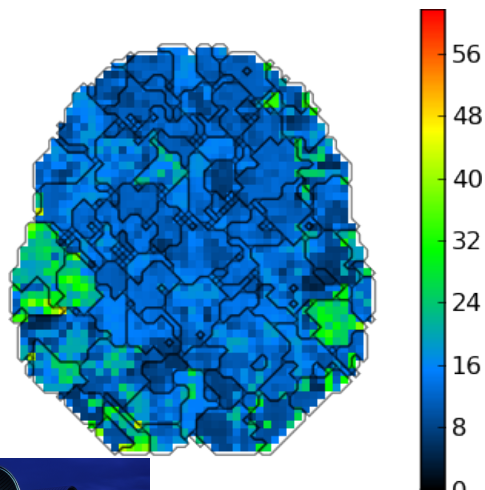
$$\hat{a}^1 - \hat{a}^2 \quad \beta=0.3$$



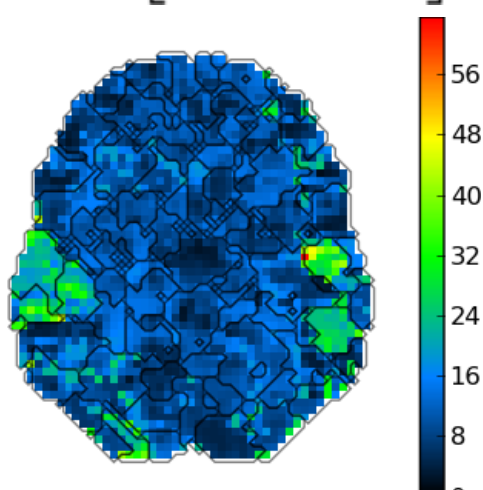
$$\hat{a}^1 - \hat{a}^2 \quad \beta=0.7$$



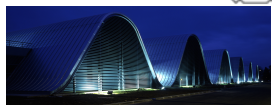
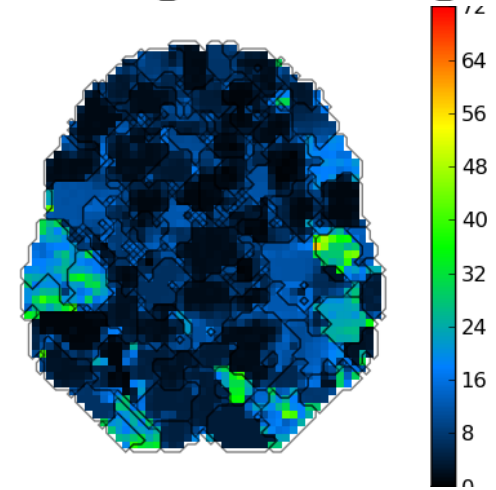
$$\text{var}[\hat{a}^1 - \hat{a}^2]$$



$$\text{var}[\hat{a}^1 - \hat{a}^2]$$



$$\text{var}[\hat{a}^1 - \hat{a}^2]$$





# Adaptive spatial regularization

- RW Metropolis-Hastings step for sampling  $\beta$
- Preliminary estimation of the partition function  $Z(\beta)$

Importance sampling density:

$$\frac{Z(\beta)}{Z(\beta')} = \mathbb{E}_{\beta'} \left[ \frac{\exp(\beta \mathcal{U}(\mathbf{q}))}{\exp(\beta' \mathcal{U}(\mathbf{q}))} \right] \text{ with } p_{\beta'}(\cdot) = Z(\beta')^{-1} \exp(\beta' \mathcal{U}(\cdot))$$

Practical implementation:

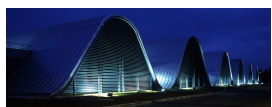
Tabulate  $Z(\beta)$  over a fine grid ( $\beta_0 = 0, \dots, \beta_D$ )

For  $d = 1, \dots, D$

- › Generate  $(\mathbf{q}_k)_{k=1:K}$  of  $p_{\beta_{d-1}}$  (SW scheme)
- › Compute  $\log Z_{\text{MCMC}}(\beta_d)$

$$\log Z_{\text{MCMC}}(\beta_d) = \log Z_{\text{MCMC}}(\beta_{d-1}) + \log \left( \frac{1}{K-I} \sum_{k=I+1}^K \exp((\beta_d - \beta_{d-1}) \mathcal{U}(\mathbf{q}_k)) \right)$$

**[Meng and Rubin, Biometrika 1998]**



# Hyper-parameter inference

- Random walk Metropolis-Hastings step for  $\beta$

$$\alpha(\beta^{(t)} \rightarrow \beta^{(c)}) = \min(1, A_{t \rightarrow c})$$

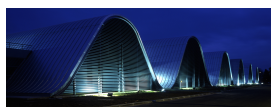
$$A_{t \rightarrow c} = \frac{p(\beta^{(c)} | \mathbf{x}^{(t)}) g(\beta^{(t)} | \beta^{(c)})}{p(\beta^{(t)} | \mathbf{x}^{(t)}) g(\beta^{(c)} | \beta^{(t)})}$$

$$= \frac{Z(\beta^{(t)})}{Z(\beta^{(c)})} \exp(-(\beta^{(c)} - \beta^{(t)})U(\mathbf{x}^{(t)})) B_{t \rightarrow c}$$

- Special case:

$$g(\cdot | x) \sim \mathcal{N}_{[0, \beta_{\max}]}(x, \xi^2) \implies B_{t \rightarrow c} = \frac{\operatorname{erf}(-\xi^{-2}\beta^{(c)}) \operatorname{erf}(\xi^{-2}(\beta_{\max} - \beta^{(t)}))}{\operatorname{erf}(-\xi^{-2}\beta^{(t)}) \operatorname{erf}(\xi^{-2}(\beta_{\max} - \beta^{(c)}))}$$

- Alternative: Gibbs sampling on the discrete grid





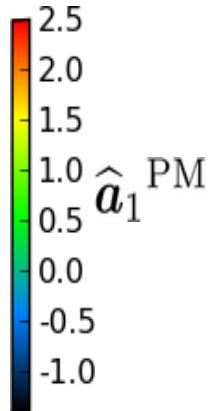
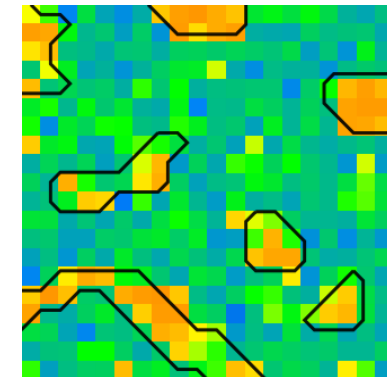
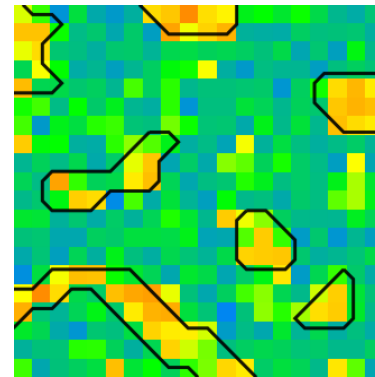
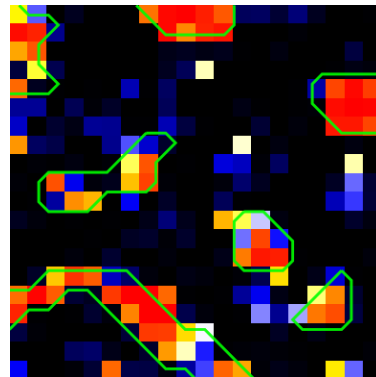
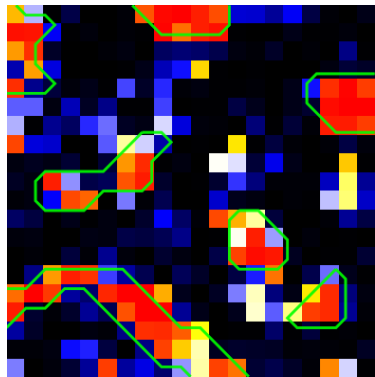
# Supervised vs. unsupervised

SMM  $\beta_1 = 0.7$

$\hat{\beta}_1^{\text{PM}} = 0.86$

SMM  $\beta_1 = 0.7$

$\hat{\beta}_1^{\text{PM}} = 0.86$

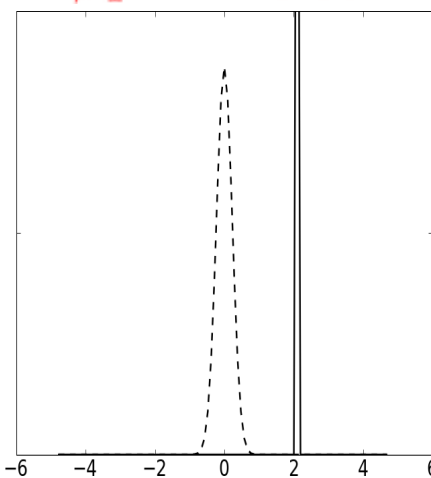
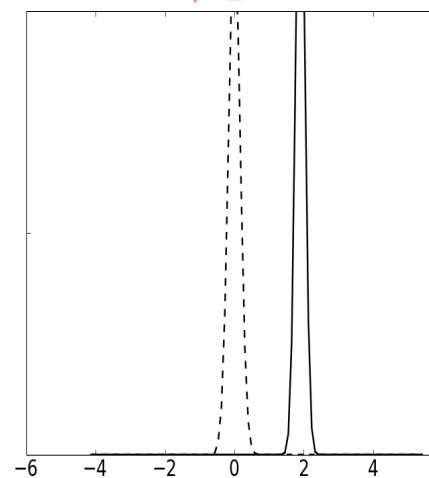
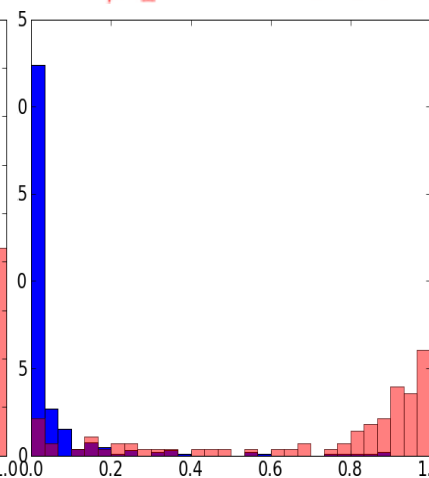
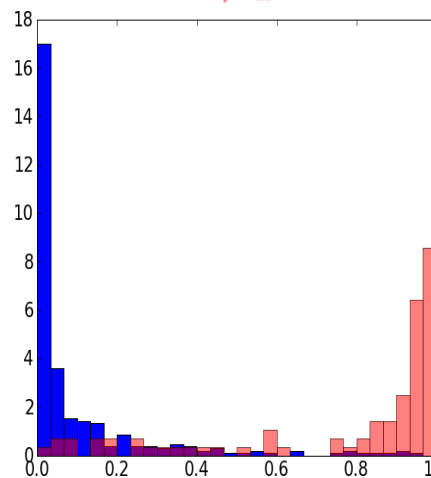


SMM  $\beta_1 = 0.7$

$\hat{\beta}_1^{\text{PM}} = 0.86$

SMM  $\beta_1 = 0.7$

$\hat{\beta}_1^{\text{PM}} = 0.86$



$\hat{\theta}_1^{\text{PM}}$

USMM



Interest of unsupervised spatial mixture models

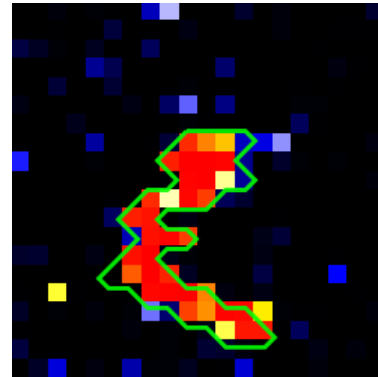
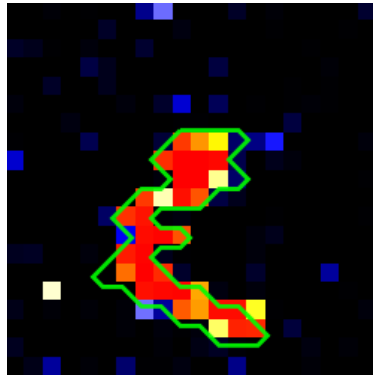




# Supervised vs. unsupervised

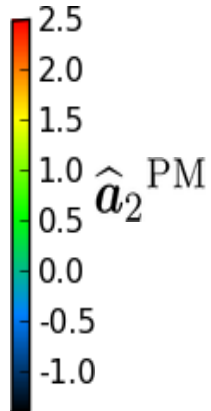
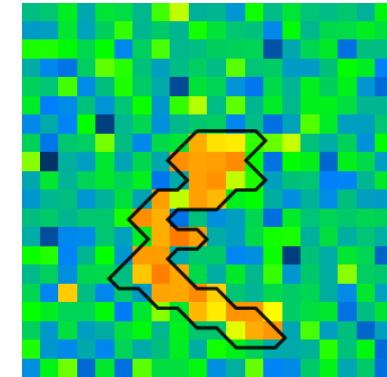
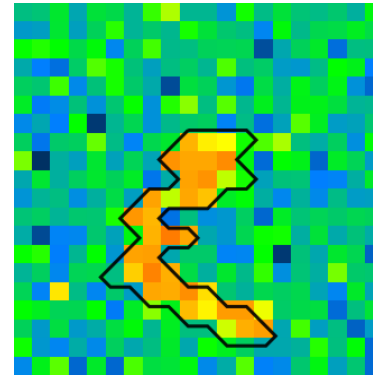
SMM  $\beta_2 = 1.1$

$\hat{\beta}_2^{\text{PM}} = 0.99$



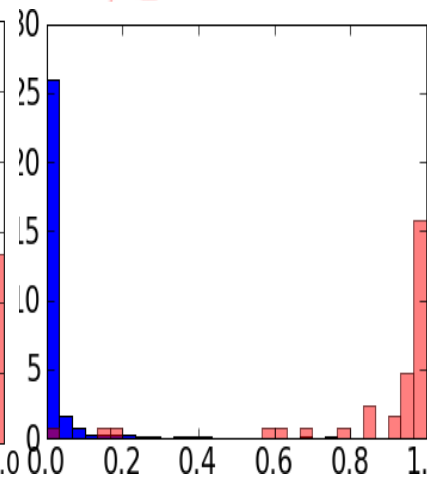
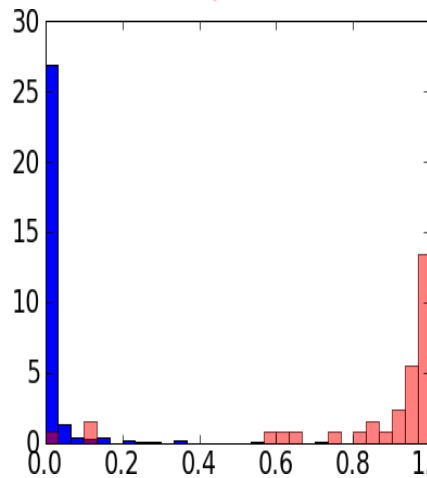
SMM  $\beta_2 = 1.1$

$\hat{\beta}_2^{\text{PM}} = 0.99$



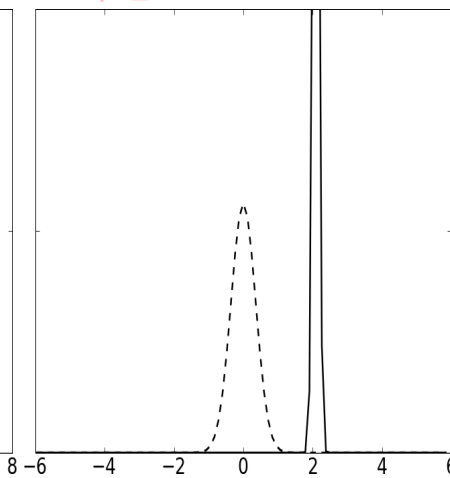
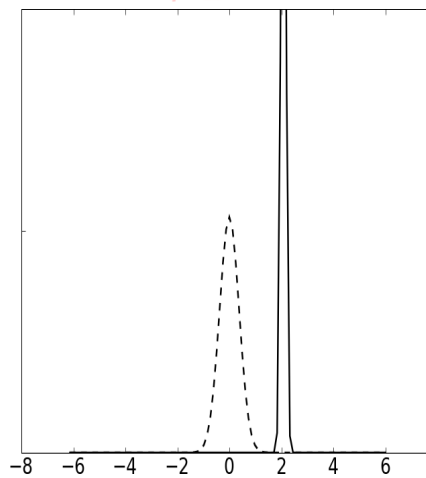
SMM  $\beta_2 = 1.1$

$\hat{\beta}_2^{\text{PM}} = 0.99$



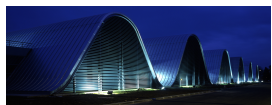
SMM  $\beta_2 = 1.1$

$\hat{\beta}_2^{\text{PM}} = 0.99$



$\hat{\theta}_2^{\text{PM}}$

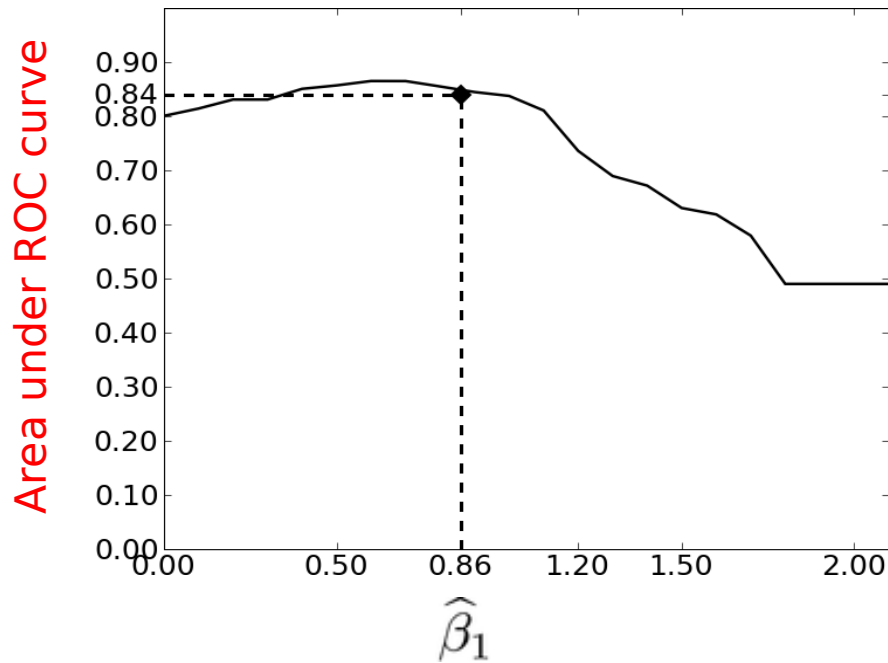
Best tuning: larger  $\beta_2$   $\longrightarrow$  lower SNR for  $m = 2$



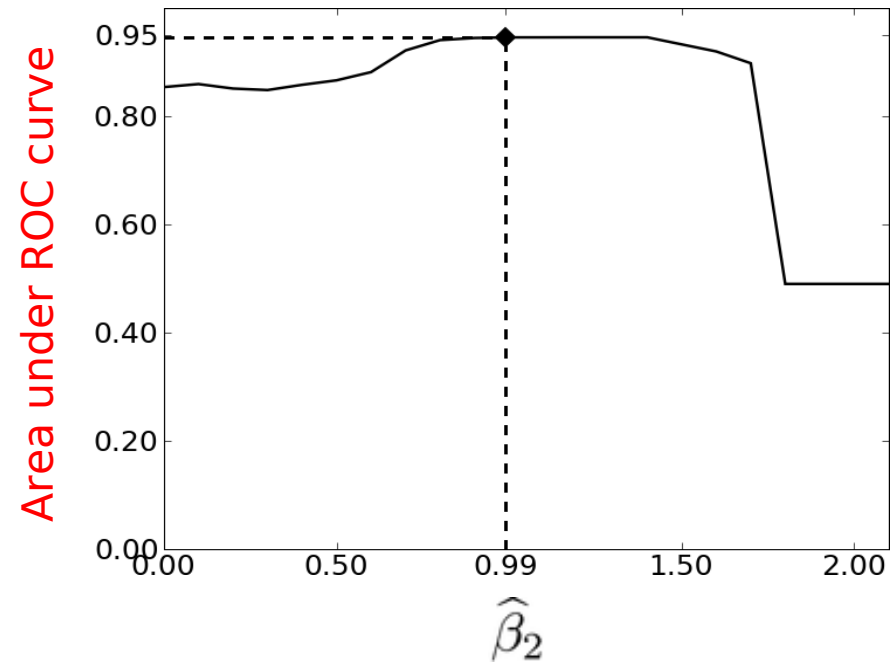


# Area under ROC curves

$m = 1$



$m = 2$



➔ nearly optimal unsupervised settings



# Remarks

- Starting point for path sampling:

$$Z(0) = 2^J \text{ for Ising fields}$$

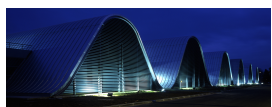
$$Z(0) = C^J \text{ for Potts fields with } C \text{ states}$$

- PF dependence on topological configuration

$$Z(\beta) = f(J, C) \text{ where } C = |\mathcal{C}| = \text{number of cliques}$$

- Path sampling for multiple data defined on the same grid
  - Which solution to adopt in other situations?
- Numerical cost of path sampling:
    - fluctuates with  $D$  (number of grid points)
    - fluctuates with MRFs  $(x^k)_{k=1:K}$  and  $\beta$  value
    - depends on MRF sampling algorithm (Gibbs, SW, PD)

[Higdon, JASA 1998]



# Adaptive spatial regularization

- Remarks:

$$Z(0) = 2^J \text{ for Ising fields}$$

$$Z(0) = C^J \text{ for Potts fields with } C \text{ states}$$

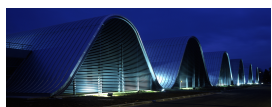
$$Z(\beta^m) = Z(\beta^n) \quad \forall n \neq m$$

$$Z(\beta) = f(J, C) \text{ where } C = |\mathcal{C}| = \text{number of cliques}$$



Extrapolation methods required for ROI of variable size and shape

**Parcel-dependent** regularization factor  $\beta$



# Linear extrapolation of $Z(\beta)$

- Linear interpolation technique: [Trillon and Idier, Eusipco 2008]

- Reference grids:

$$(\mathcal{G}_p)_{p=1:P} \implies (\log \hat{Z}_{\mathcal{G}_p}(\beta_k))_{p=1:P}, \forall \beta_k = k\Delta\beta$$

- Linear regression:

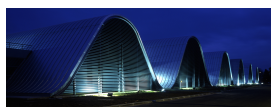
$$\forall \beta_k, (A_{\beta_k}, B_{\beta_k}) = \arg \min_{(A,B) \in \mathbb{R}^2} \sum_{p=1}^P \|\log \hat{Z}_{\mathcal{G}_p}(\beta_k) - Ac_p - B\|^2$$

- Application of linear interpolation to test grid:

$$\forall \beta_k, \log \tilde{Z}_{\mathcal{T}}(\beta_k) = A_{\beta_k} c_{\mathcal{T}} + B_{\beta_k}$$



Require homogeneity of the reference grids and regular grids !!



# Bilinear extrapolation of $Z(\beta)$

- **Bilinear extension:** [Risser, Idier, Ciuciu, ICIP 2009]

- Bilinear regression:

$$\forall \beta_k, (A_{\beta_k}, B_{\beta_k}, D_{\beta_k}) = \arg \min_{(A, B, D) \in \mathbb{R}^3} \sum_{p=1}^P \|\log \hat{Z}_{G_p}(\beta_k) - A c_p - B s_p - D\|^2$$

- Application of bilinear interpolation to test grid:

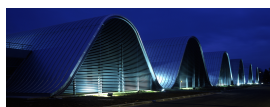
$$\forall \beta_k, \quad \log \tilde{Z}_{\mathcal{T}}(\beta_k) = A_{\beta_k} c_{\mathcal{T}} + B_{\beta_k} s_{\mathcal{T}} + D_{\beta_k}$$

Still homogeneous reference set but applicable to



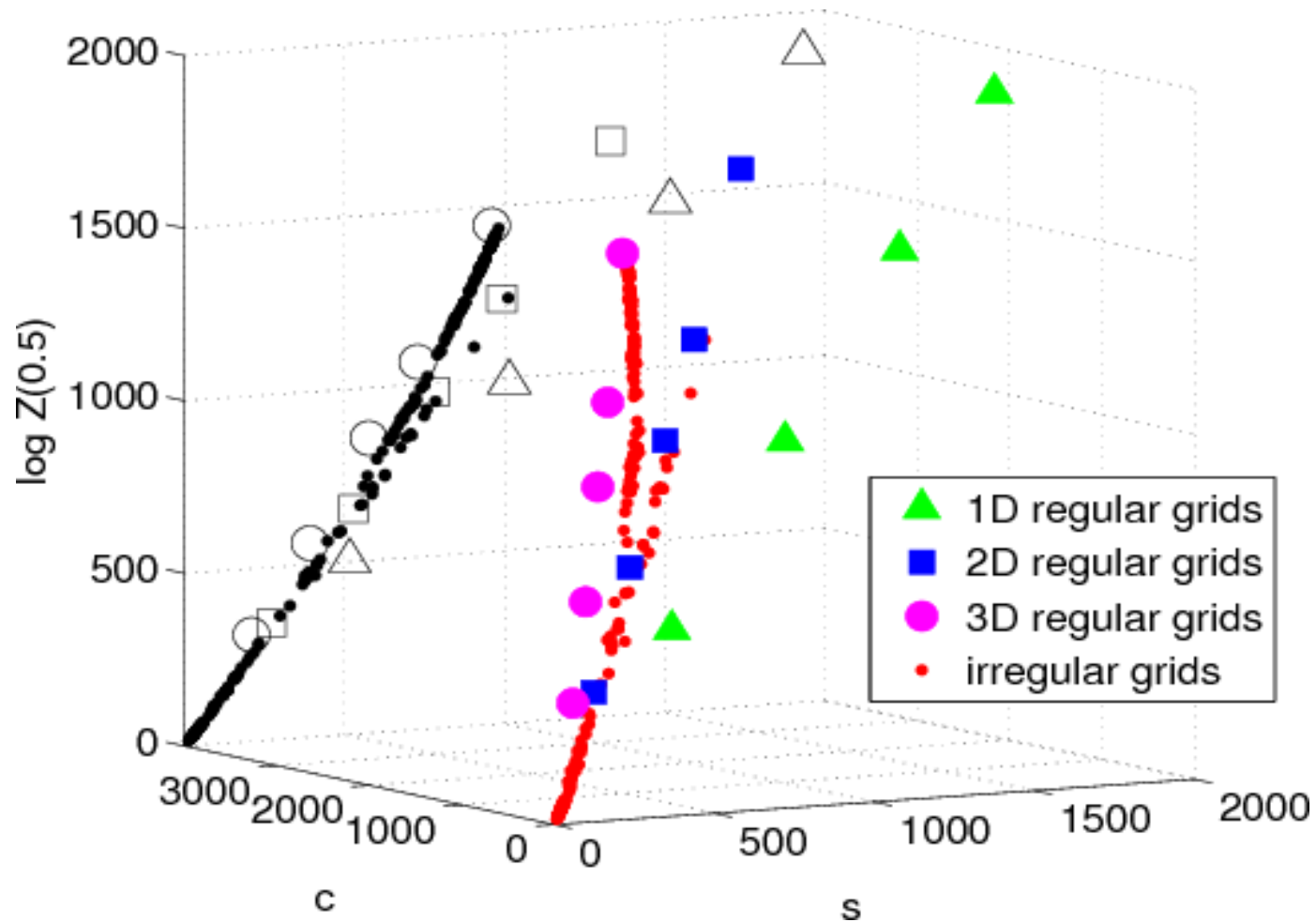
**non regular grids**

**Multiple PFs** involved in the extrapolation process

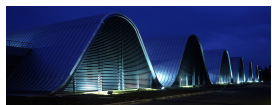


# Bilinear dependence of $Z(\beta)$

- Illustration



At a fixed number of  $C$ , the larger  $S$  the larger  $\log Z(\beta)$



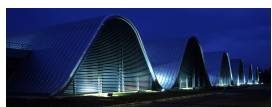


# Comparison of linear/bilinear

Mean approximation error over **Regular** & **Irregular** test fields.  
Errors given in percentage.

Test grid		Scheme / Reference grid			
		B=bilinear, L=linear / R=regular, I=irregular			
		B / R	B / I	L / R	L / I
regular	small	<b>0.747</b>	3.84	5.55	93.0
	medium	1.30	<b>0.991</b>	7.27	6.37
	large	1.59	<b>1.31</b>	9.18	7.18
irregular	$\beta = 0.2$	6.85	<b>1.29</b>	23.6	83.9
	$\beta = 0.4$	0.984	<b>0.264</b>	7.71	8.28
	$\beta = 0.5$	1.73	<b>1.27</b>	1.64	1.52

Improved performance with the bilinear approach for **small and irregular fields**  
Approximation accuracy depends on  $\beta$



# Min/max extrapolation of $Z(\beta)$

- Fast extrapolation technique: [\[Risser, Ciuciu et al, MLSP 2009\]](#)  
[\[Risser, Ciuciu et al, MICCAI 2009\]](#)

- Reference grids:

$$(\mathcal{G}_p)_{p=1:P} \implies (\log \hat{Z}_{\mathcal{G}_p}(\beta_k))_{p=1:P}, \forall \beta_k = k\Delta\beta$$

- Grid selection: Min/max criterion

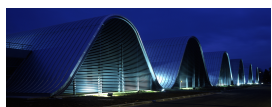
$$\mathcal{G}_{\text{ref}} = \arg \min_{(\mathcal{G}_p)_{p=1:P}} \mathcal{A}_{\mathcal{T}}(0, \mathcal{G}_p) \text{ with } \mathcal{A}_{\mathcal{T}}(\beta, \mathcal{G}_p) = \|\log \hat{Z}_{\mathcal{T}}(\beta) - \log \tilde{Z}_{\mathcal{T}}(\beta)\|^2$$

- Extrapolation:  $\log \tilde{Z}_{\mathcal{T}}(\beta) = \frac{c_{\mathcal{T}}}{c_{\mathcal{G}_{\text{ref}}}} \left( \log \hat{Z}_{\mathcal{G}_{\text{ref}}}(\beta) - \log L \right) + \log L$

- **Maximal** error:  $\mathcal{A}_{\mathcal{T}}(0, \mathcal{G}_p) = \|\log L \left[ (s_{\mathcal{T}} - 1) - \frac{c_{\mathcal{T}}}{c_{\mathcal{G}_p}} (s_{\mathcal{G}_p} - 1) \right]\|^2$

**Single PF estimate** involved in the extrapolation

The more **different** the reference grids **the smaller** the approximation error





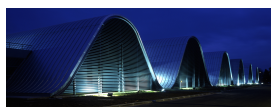
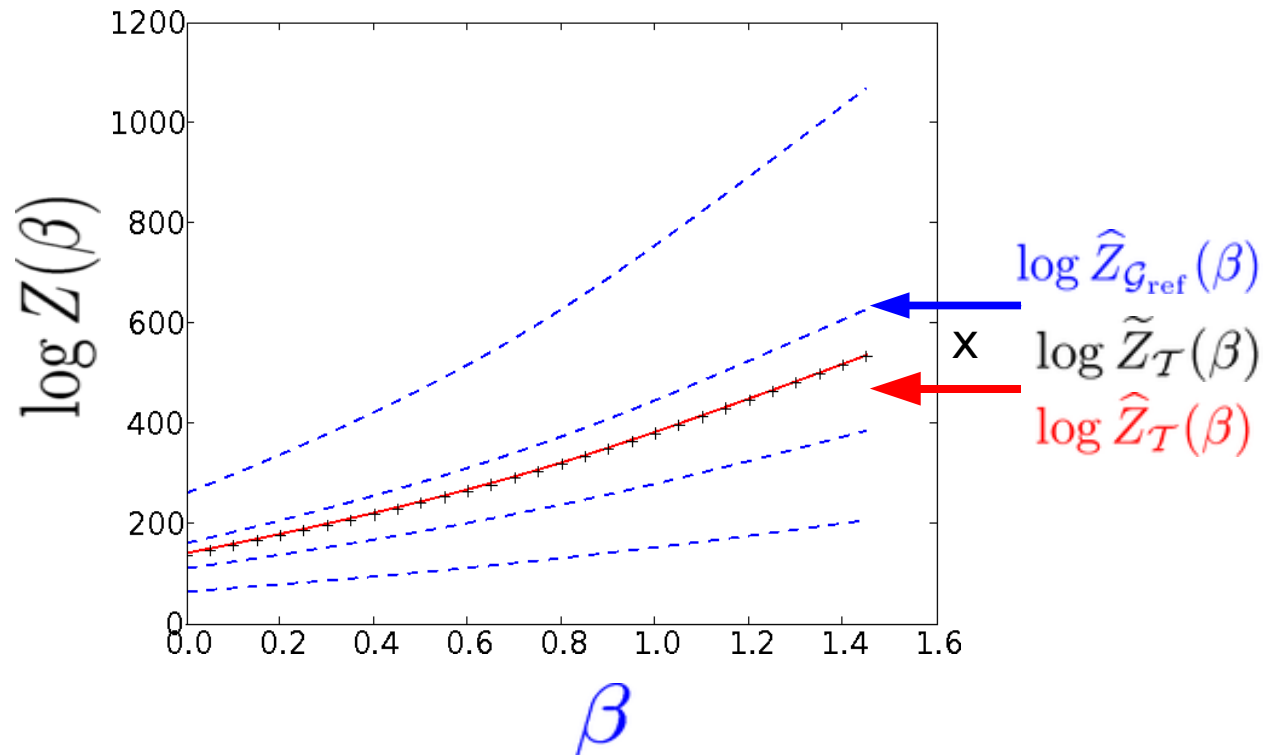
# Min/max extrapolation of $Z(\beta)$

- Fast extrapolation technique:

- Grid selection:

$$\mathcal{G}_{\text{ref}} = \arg \min_{(\mathcal{G}_p)_{p=1:P}} \mathcal{A}_{\mathcal{T}}(0, \mathcal{G}_p) \text{ with } \mathcal{A}_{\mathcal{T}}(\beta, \mathcal{G}_p) = \|\log Z_{\mathcal{T}}(\beta) - \log \tilde{Z}_{\mathcal{T}}(\beta)\|^2$$

- Maximum error:  $\mathcal{A}_{\mathcal{T}}(0, \mathcal{G}_p) = \|\log L[(s_{\mathcal{T}} - 1) - \frac{c_{\mathcal{T}}}{c_{\mathcal{G}_p}}(s_{\mathcal{G}_p} - 1)]\|^2$



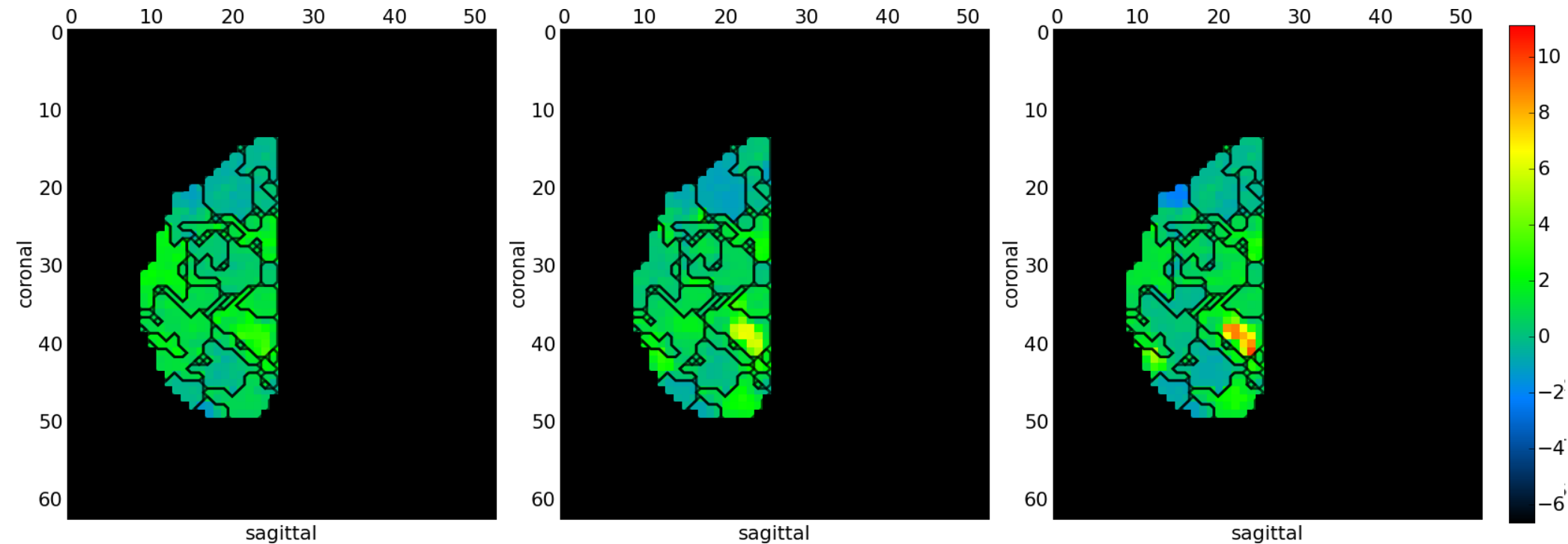
# Half whole brain analysis

Normalized contrast: Audit (Calculation - Sentence)

IMM

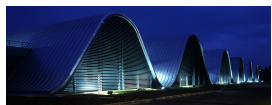
SSMM  $\beta=0.8$

USMM



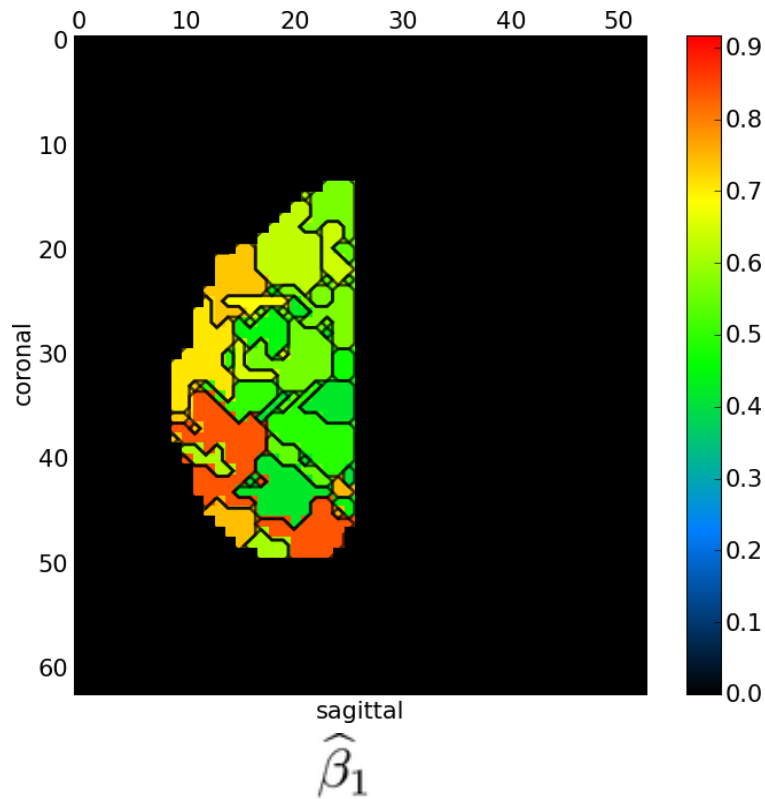
Activations enhanced in the parietal cortex using U/SSMM  
Coherence with sulcal anatomy & literature

[Vincent, Ciuciu et al, in rev. IEEE TMI 2009]

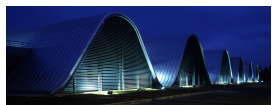
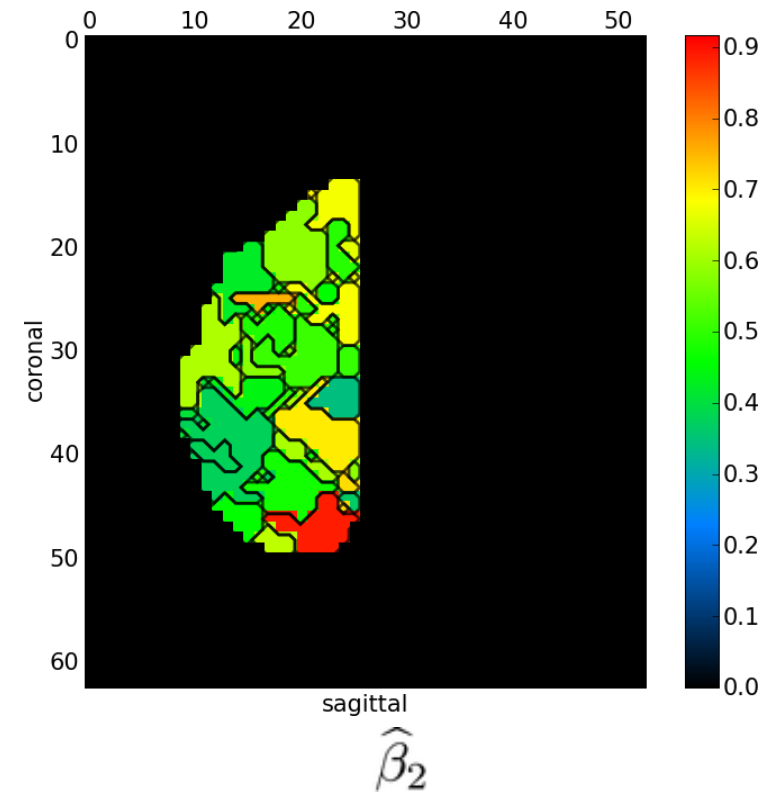


# Adaptive spatial regularization

## Auditory calculation



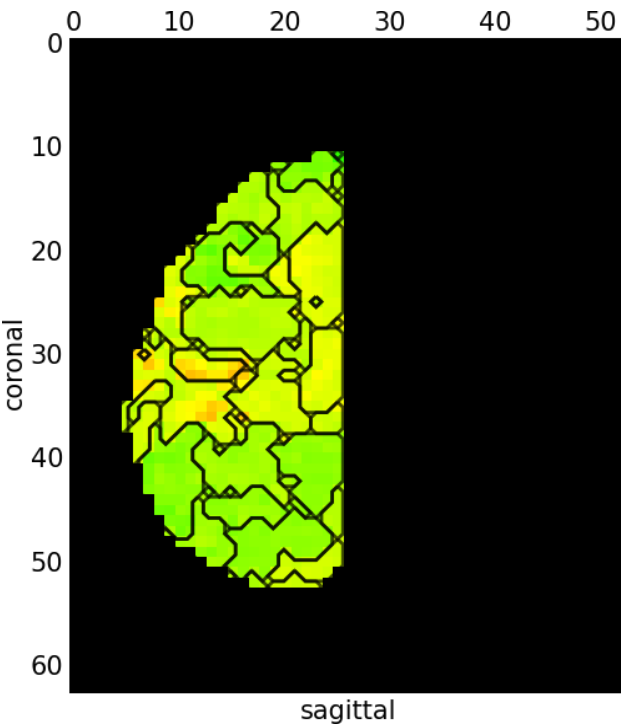
## Auditory sentence



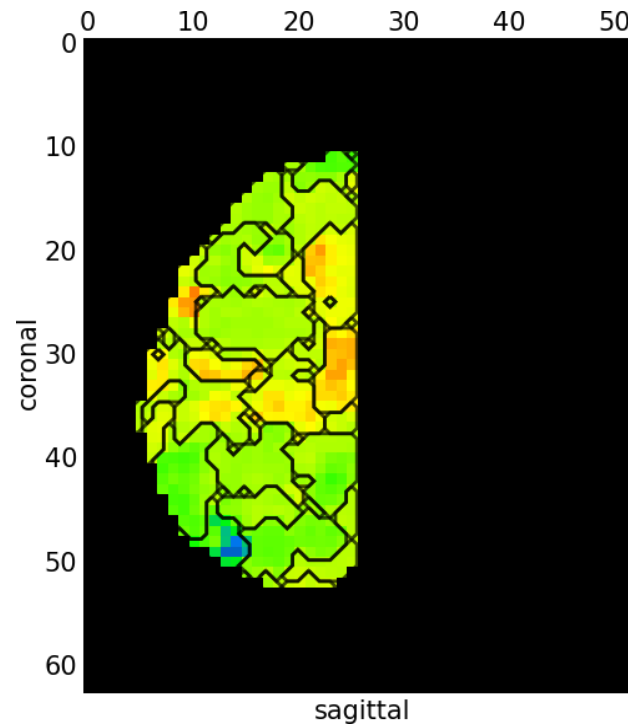
# Half whole brain analysis

Normalized contrast: Right - left “auditory clicks”

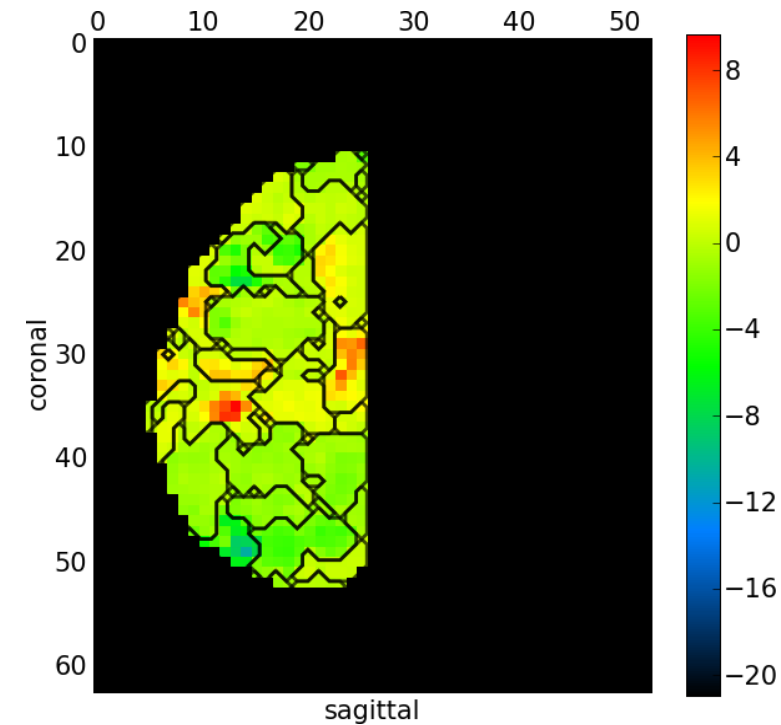
IMM



SSMM  $\beta=0.8$



USMM

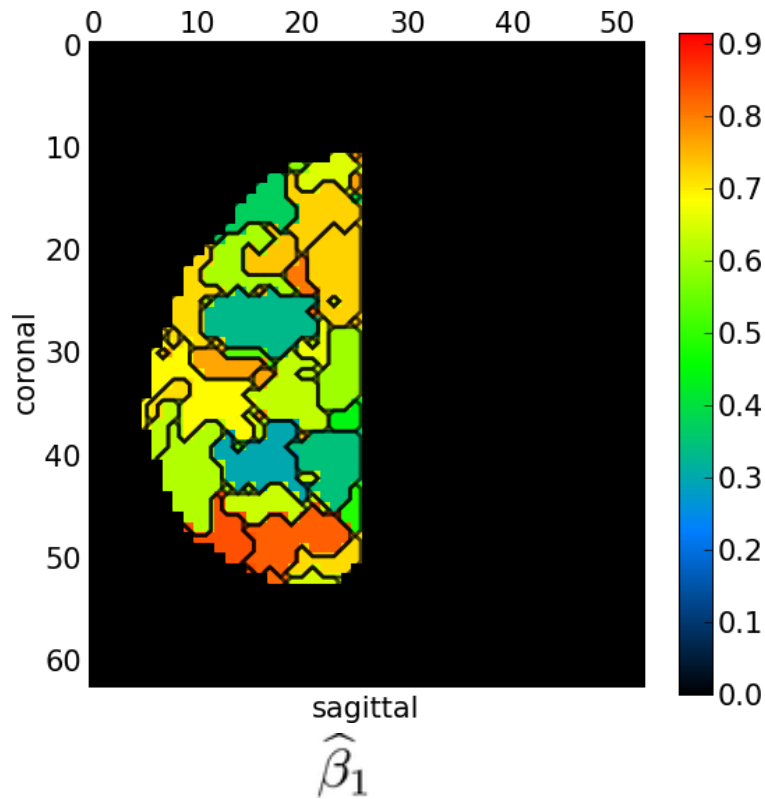


Only USMM provides more sensitive activation in the left motor cortex

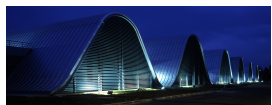
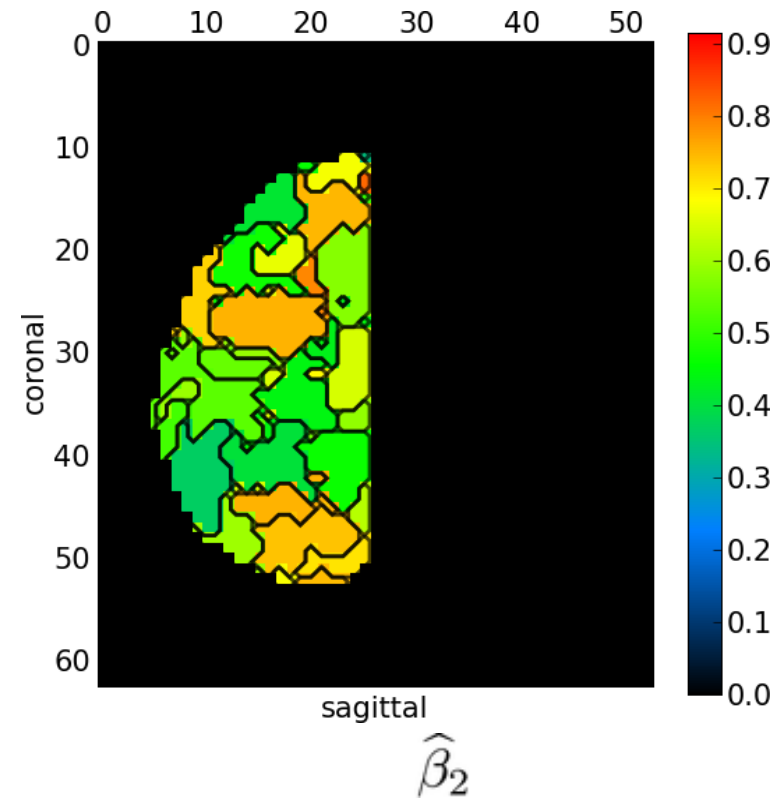


# Adaptive spatial regularization

Right click



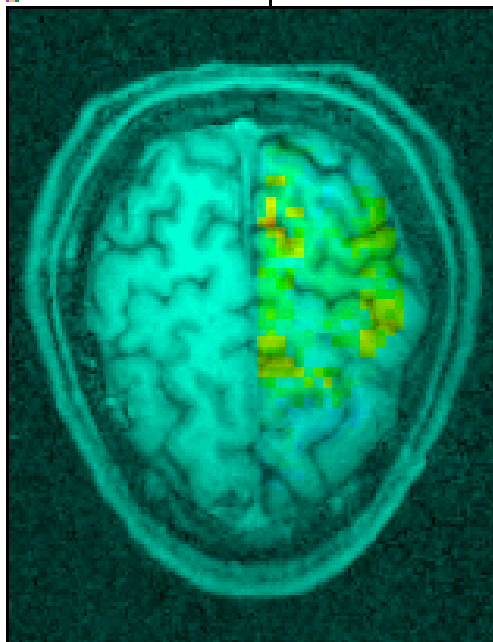
Left click



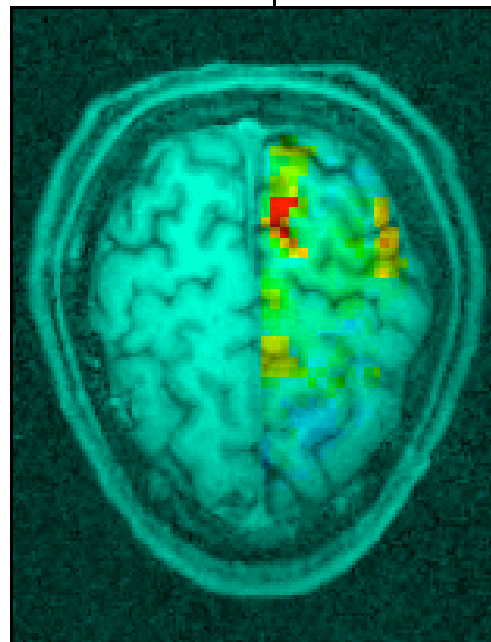
# Whole brain analysis

Normalized contrast: Audit. (Calculation - Sentence)

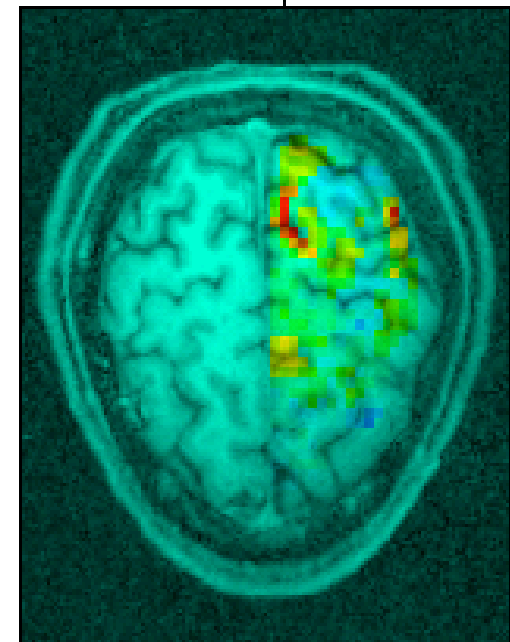
IMM



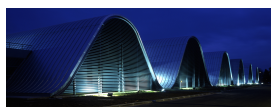
SSMM  $\beta=0.8$



USMM



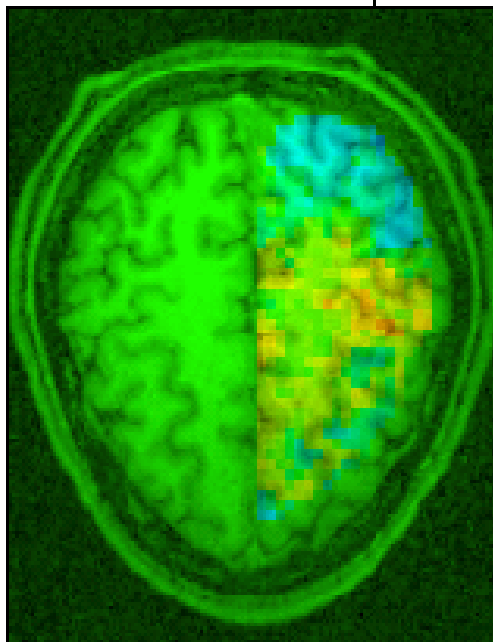
→ Activations enhanced in the parietal cortex using U/SSMM  
Coherent with sulcal anatomy



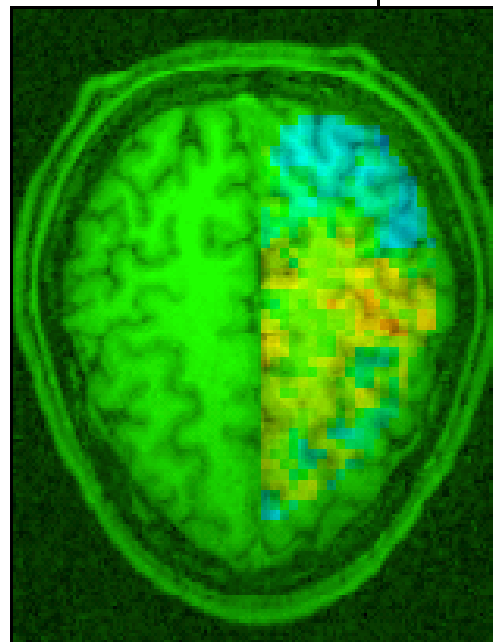
# Half whole brain analysis

Normalized contrast: Left - right "auditory clicks"

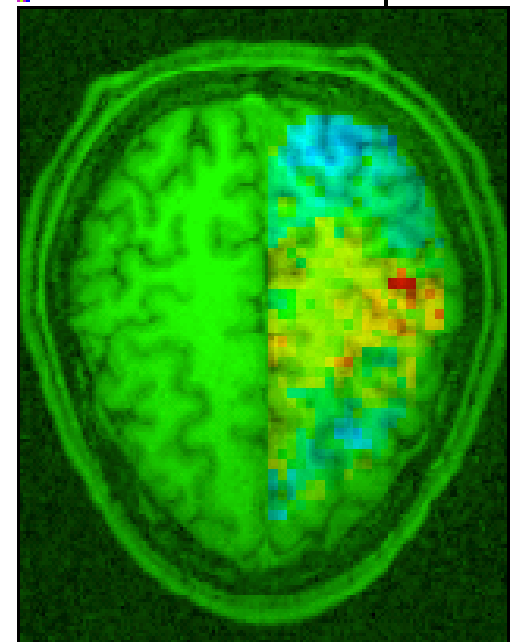
IMM



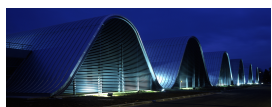
SSMM  $\beta=0.8$



USMM



Only USMM provides more sensitive activation in the motor cortex



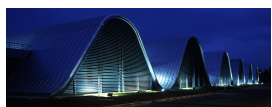
# Scale ambiguity

- **Identifiability problem**

[Veit and Idier, TS 2009]  
[Ciuciu et al, GRETSI 2007]

$$p(\mathbf{y} | \mathbf{h}, \mathbf{a}, \mathbb{1}, \boldsymbol{\theta}_0) = p(\mathbf{y} | \mathbf{h}/s, s\mathbf{a}, \mathbb{1}, \boldsymbol{\theta}_0), \forall s \neq 0$$

- A common issue to all bilinear inverse problems:
  - Blind source separation
  - Blind deconvolution
  - Joint detection-estimation
- **Bayesian inference:** proper priors help in solving this ambiguity





# Gibbs sampling

1.  $\mathbf{a}^{(k+1)} \leftarrow \mathbf{A} \sim f_{\mathbf{A}|\mathbf{H},\mathbf{Y},\Theta}(\mathbf{a} | \mathbf{h}^{(k)}, \mathbf{y}, \boldsymbol{\theta}^{(k)})$
2.  $\mathbf{h}^{(k+1)} \leftarrow \mathbf{H} \sim f_{\mathbf{H}|\mathbf{A},\mathbf{Y},\Theta}(\mathbf{h} | \mathbf{a}^{(k+1)}, \mathbf{y}, \boldsymbol{\theta}^{(k)})$
3.  $\boldsymbol{\theta}^{(k+1)} \leftarrow \Theta \sim f_{\Theta|\mathbf{A},\mathbf{H},\mathbf{Y}}(\boldsymbol{\theta} | \mathbf{a}^{(k+1)}, \mathbf{h}^{(k+1)}, \mathbf{y})$

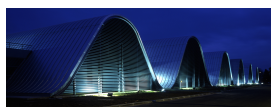
- Slow evolution of the scale



Slow convergence of Gibbs sampling  
Wrong mixing properties of the MC

- **Solutions available in the literature:**

- Do nothing: assume an implicit shape/scale decorrelation
- Normalize at a fixed arbitrary scale at each iteration  
(deterministic transformation incompatible with target density)



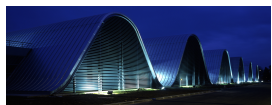
# Scale sampling

- Introduce an additional step in Gibbs sampling

$$\mathbf{A} = \mathbf{A}_{\text{old}} \times S, \quad \mathbf{H} = \mathbf{H}_{\text{old}}/S$$

- Up to now:  $S = \|\mathbf{H}\|$
- Alternative:  $S$  is a random variable to be sampled

according to which pdf?



# Scale sampling (cont'd)

- General principle: **make a change of variable**

$$\begin{aligned}
 (s, \mathbf{v}) &= \phi(\mathbf{a}, \mathbf{h}) \\
 &= \left( \underbrace{a_1/a_1^{\text{old}}}_s, \underbrace{a_2/a_1, \dots, a_J/a_1}_{v_1, \dots, v_{J-1}}, \underbrace{a_1 h_1, \dots, a_1 h_P}_{v_J, \dots, v_{J+P-1}} \right).
 \end{aligned}$$

$$f_{S, \mathbf{V} | \Theta}(s) \propto |s|^{J-P-1} f_{\mathbf{A}, \mathbf{H} | \Theta}(\phi^{-1}(s, \mathbf{v}))$$



# Scale sampling (cont'd)

- General principle: **make a change of variable**

$$\begin{aligned}
 (s, \mathbf{v}) &= \phi(\mathbf{a}, \mathbf{h}) \\
 &= \left( \underbrace{a_1/a_1^{\text{old}}}_s, \underbrace{a_2/a_1, \dots, a_J/a_1}_{v_1, \dots, v_{J-1}}, \underbrace{a_1 h_1, \dots, a_1 h_P}_{v_J, \dots, v_{J+P-1}} \right).
 \end{aligned}$$

$$f_{S|\mathbf{V},\Theta}(s) \propto f_{S,\mathbf{V}|\Theta}(s) \propto |s|^{J-P-1} f_{\mathbf{A},\mathbf{H}|\Theta}(\phi^{-1}(s, \mathbf{v}))$$



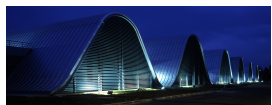
# Scale sampling (cont'd)

- General principle: **make a change of variable**

$$\begin{aligned}
 (s, \mathbf{v}) &= \phi(\mathbf{a}, \mathbf{h}) \\
 &= \left( \underbrace{a_1/a_1^{\text{old}}}_s, \underbrace{a_2/a_1, \dots, a_J/a_1}_{v_1, \dots, v_{J-1}}, \underbrace{a_1 h_1, \dots, a_1 h_P}_{v_J, \dots, v_{J+P-1}} \right).
 \end{aligned}$$

- $S$  is independent of the data  $\mathbf{y}$

$$f_{S|\mathbf{V}, \mathbf{Y}, \Theta}(s) = f_{S|\mathbf{V}, \Theta}(s) \propto f_{S, \mathbf{V}|\Theta}(s) \propto |s|^{J-P-1} f_{\mathbf{A}, \mathbf{H}|\Theta}(\phi^{-1}(s, \mathbf{v}))$$

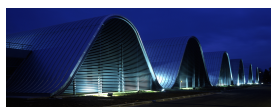




# Modified Gibbs sampling

Given  $\mathbf{A}^{(k)}, \mathbf{h}^{(k)}, \boldsymbol{\theta}^{(k)}$

1.  $\mathbf{a}_{\text{old}}^{(k+1)} \leftarrow \mathbf{A} \sim f_{\mathbf{A}|\mathbf{H},\mathbf{Y},\boldsymbol{\Theta}}(\mathbf{a} | \mathbf{h}^{(k)}, \mathbf{y}, \boldsymbol{\theta}^{(k)})$
2.  $\mathbf{h}_{\text{old}}^{(k+1)} \leftarrow \mathbf{H} \sim f_{\mathbf{H}|\mathbf{A},\mathbf{Y},\boldsymbol{\Theta}}(\mathbf{h} | \mathbf{a}_{\text{old}}^{(k+1)}, \mathbf{y}, \boldsymbol{\theta}^{(k)})$
3.  $s \leftarrow S \sim f_{S|\text{rest}}(s) \propto |s|^{MJ-P-1} f_{\mathbf{A},\mathbf{H}|\boldsymbol{\Theta}}(s\mathbf{a}_{\text{old}}^{(k+1)}, \mathbf{h}_{\text{old}}^{(k+1)} / s | \boldsymbol{\theta})$   
 $\mathbf{a}^{(k+1)} \leftarrow s\mathbf{a}_{\text{old}}^{(k+1)}$  and  $\mathbf{h}^{(k+1)} \leftarrow \mathbf{h}_{\text{old}}^{(k+1)} / s$
4.  $\boldsymbol{\theta}^{(k+1)} \leftarrow \boldsymbol{\Theta} \sim f_{\boldsymbol{\Theta}|\mathbf{A},\mathbf{H},\mathbf{Y}}(\boldsymbol{\theta} | \mathbf{a}^{(k+1)}, \mathbf{h}^{(k+1)}, \mathbf{y})$

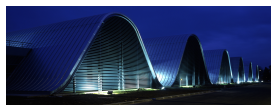


# Examples

- Gaussian priors on  $(\mathbf{H}, \mathbb{A}) \longrightarrow f_{S^2 | rest} \sim GIG(\lambda, \alpha, \beta)$
- Gamma priors on  $(\mathbf{H}, \mathbb{A}) \longrightarrow f_{S | rest} \sim GIG(\lambda, \alpha, \beta)$
- Still valid for Gaussian mixtures on  $\mathbb{A}$
- JDE framework:

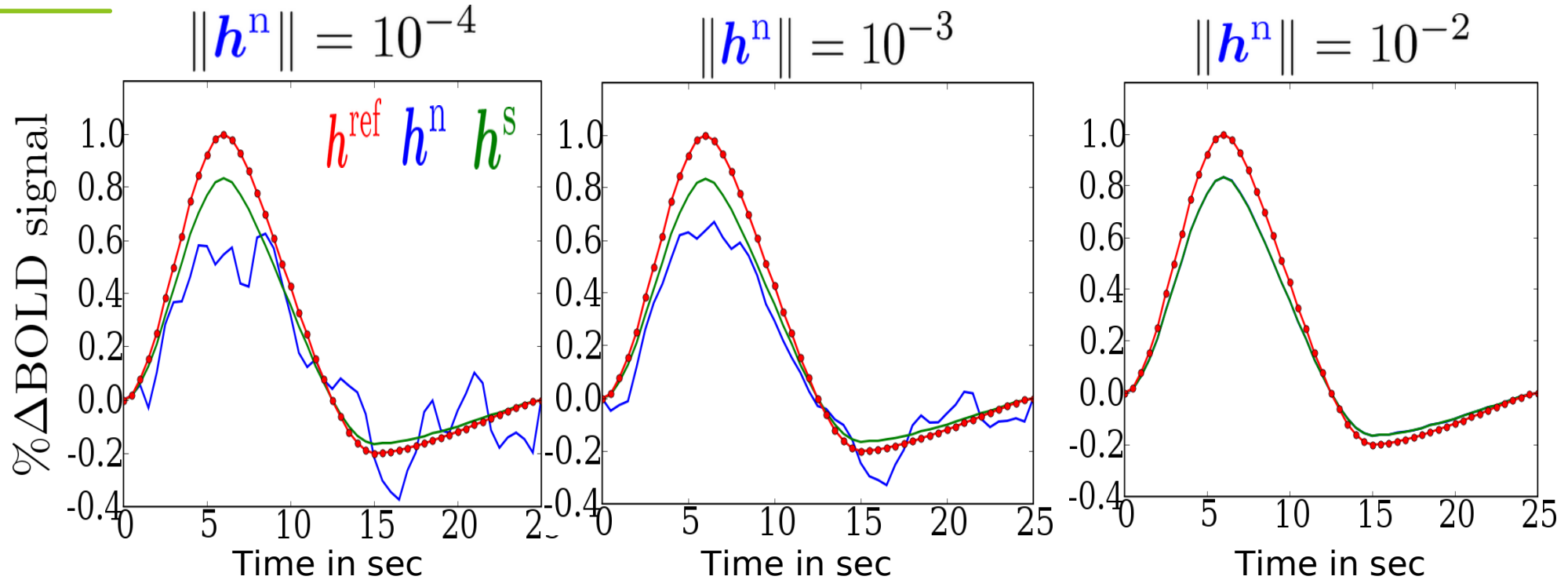
$$\lambda = (P - M(J + 1))/2, \quad \alpha = (\mathbf{h}_{old})^t \mathbf{R}^{-1} \mathbf{h}_{old} / \sigma_h^2,$$

$$\beta = \frac{\|\boldsymbol{\mu}_{old}\|^2}{2\sigma_\mu^2} + \sum_{j=1}^J ((\mathbf{a}_j)_{old})^t \Sigma_j^{-1} (\mathbf{a}_j)_{old} \text{ with } \Sigma_j = \text{diag}_M[v_{q_j^m}^m]$$

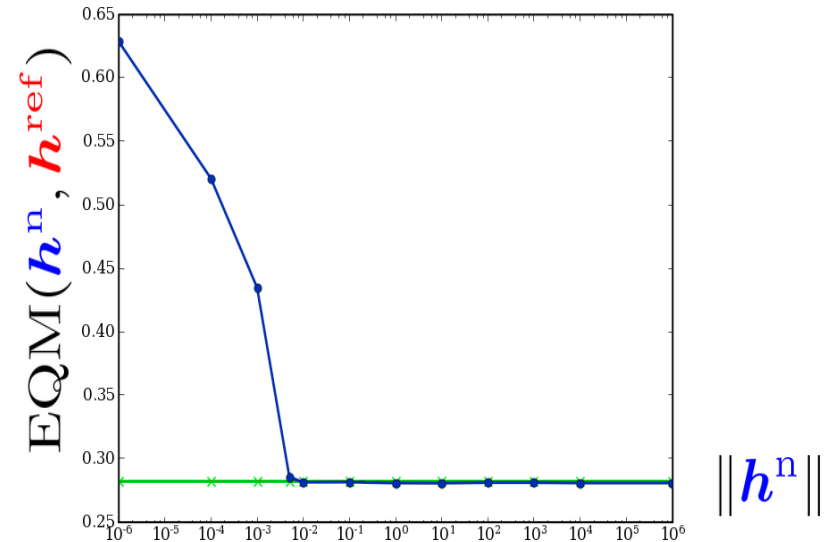




# Illustrations



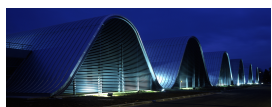
$$\text{EQM}(h^n, h^{\text{ref}}) = \|h^n - h^{\text{ref}}\|^2$$





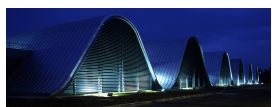
# Conclusions

- The joint detection-estimation framework:
  - directly accounts for different sources of variability
  - provides both region-based HRF time courses and contrast maps
  - embeds unsupervised spatial regularization
  - avoids using spatial filtering of fMRI datasets
  - depends on an input parcellation: **[Vincent et al, ISBI'08]**
  - gives improved RFX maps in comparison with SPM
- First release of Pyhrf package (v 1.0)
  - downloadable at <http://launchpad.net> (nipy project)

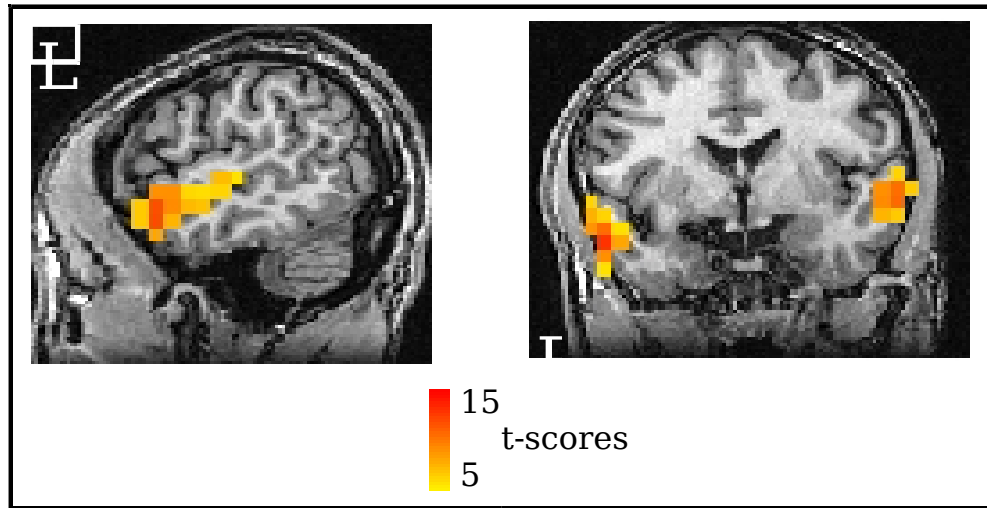


# Ongoing works

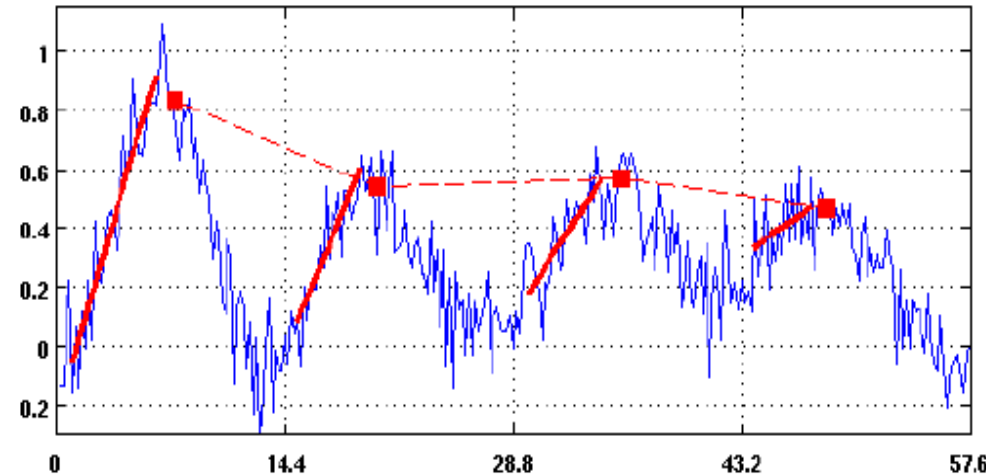
- Neuro-dynamics models (habituation effect)
- Validation at the group level
- Model comparison and selection
- Extension to the cortical surface for EEG/fMRI fusion
- Application to neonate fMRI datasets
  - collab: G. Dehaene (INSERM U562)



# Forward model with habituation

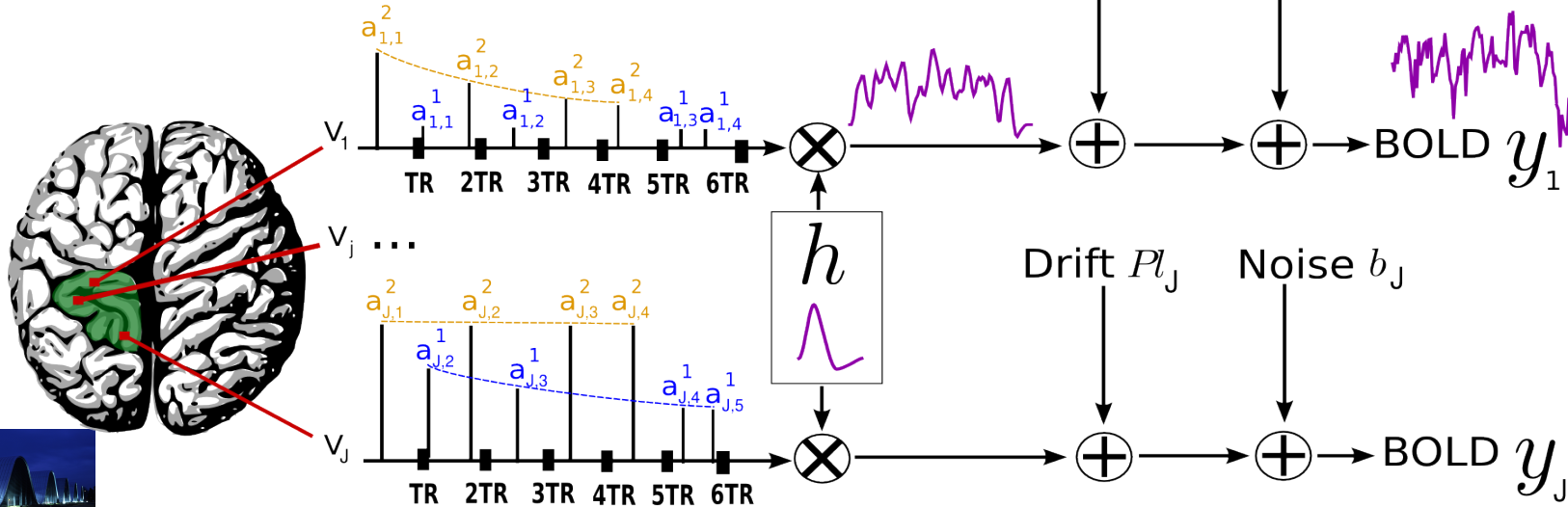


[Rabrait, Ciuciu et al, JMRI 2008]



Activation in response to the first sentence detected in the STS/G  
[Ciuciu et al, ICASSP 2009]

$a_j^m$ : neural response level for voxel  $j$  and condition  $m$



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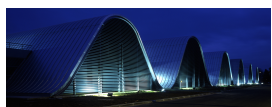
Marion Soumoy

Stéphane Sockeel

Jérôme Idier

Sophie Donnet

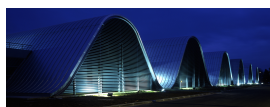
Thomas Veit



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